Detection and localization of defects for sensor array imaging in noisy environments

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Consider a simple imaging problem: detection and localization of a point reflector. \rightarrow "All" methods work.

Add noise.

 \rightarrow Which method is best ?

The array response matrix (in an ideal world)

- Time-harmonic waves emitted by point sources and recorded by point sensors
- Array of n elements $\{y_1, \ldots, y_n\}$.
- û(y_j, y_l) = field recorded by the sensor at y_j when the sensor at y_l emits a time-harmonic signal at frequency ω.



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• Response matrix $\mathbf{A}_0 = ((A_0)_{jl})_{j,l=1,...,n}$ in a known environment:

 $(A_0)_{jl} = \hat{u}(\boldsymbol{y}_j, \boldsymbol{y}_l) - \hat{G}_0(\boldsymbol{y}_j, \boldsymbol{y}_l)$

where

- $\hat{u}(y_j, y_l)$ the field recorded by the sensor at y_j when the sensor at y_l emits.
- $\hat{G}_0(\boldsymbol{y}_j, \boldsymbol{y}_l)$ is the incident field (Green's function of the background medium).

A simple model: a point reflector in a constant background (1/2)

Scalar wave equation:
$$\Delta_{\boldsymbol{x}} \hat{u}(\boldsymbol{x}, \boldsymbol{y}) + \frac{\omega^2}{c^2(\boldsymbol{x})} \hat{u}(\boldsymbol{x}, \boldsymbol{y}) = -\delta_{\boldsymbol{y}}(\boldsymbol{x}).$$

In the presence of a localized reflector in the search domain Ω :

$$\frac{1}{c^2(\boldsymbol{x})} = \frac{1}{c_0^2} (1 + \sigma_{\mathrm{r}} V_{\mathrm{ref}}(\boldsymbol{x}))$$

- c_0 is the known background speed (supposed constant),

- the local variation $V_{\text{ref}}(\boldsymbol{x}) = \mathbf{1}_{\Omega_{\text{ref}}}(\boldsymbol{x} - \boldsymbol{x}_{\text{ref}})$ represents the reflector at $\boldsymbol{x}_{\text{ref}}$, where Ω_{ref} is a compactly supported domain with volume l_{ref}^3 .

• Response matrix $\mathbf{A}_0 = ((A_0)_{jl})_{j,l=1,\ldots,n}$:

$$(A_0)_{jl} = \hat{u}(\boldsymbol{y}_j, \boldsymbol{y}_l) - \hat{G}_0(\boldsymbol{y}_j, \boldsymbol{y}_l)$$

- $\hat{u}(\boldsymbol{y}_j, \boldsymbol{y}_l)$ the field recorded by the sensor at \boldsymbol{y}_j when the sensor at \boldsymbol{y}_l emits:

$$\Delta_{\boldsymbol{x}}\hat{u}(\boldsymbol{x},\boldsymbol{y}_l) + \frac{\omega^2}{c_0^2} \big(1 + \sigma_{\rm r} V_{\rm ref}(\boldsymbol{x})\big)\hat{u}(\boldsymbol{x},\boldsymbol{y}_l) = -\delta_{\boldsymbol{y}_l}(\boldsymbol{x})$$

- $\hat{G}_0(\boldsymbol{y}_j, \boldsymbol{y}_l)$ is the incident field:

$$\hat{G}_0(\boldsymbol{y}_j, \boldsymbol{y}_l) = \frac{e^{i\frac{\omega}{c_0}|\boldsymbol{y}_j - \boldsymbol{y}_l|}}{4\pi|\boldsymbol{y}_j - \boldsymbol{y}_l|}$$

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A simple model: a point reflector in a constant background (2/2)

• If the reflector is small, the response matrix has the form (Born approximation):

$$(A_0)_{jl} = \hat{G}_0(\boldsymbol{x}_j, \boldsymbol{x}_{\mathrm{ref}}) rac{\omega^2}{c_0^2} \sigma_{\mathrm{r}} l_{\mathrm{ref}}^3 \hat{G}_0(\boldsymbol{x}_{\mathrm{ref}}, \boldsymbol{x}_l)$$

or equivalently:

$$\mathbf{A}_0 = \sigma_{ ext{ref}} oldsymbol{g}(oldsymbol{x}_{ ext{ref}})^T, ext{ with } \quad \sigma_{ ext{ref}} = rac{\omega^2}{c_0^2} \sigma_{ ext{r}} l_{ ext{ref}}^3 \Big(\sum_{l=1}^n |\hat{G}_0(oldsymbol{x}_{ ext{ref}},oldsymbol{x}_l)|^2\Big)$$

- $\sigma_{\rm r}$ is the scattering amplitude of the reflector,
- $l_{\rm ref}^3$ is the volume of the reflector,
- $\boldsymbol{g}(\boldsymbol{x})$ is the normalized vector of Green's functions from the array to the point \boldsymbol{x} :

$$\boldsymbol{g}(\boldsymbol{x}) = \frac{1}{\left(\sum_{l=1}^{n} |\hat{G}_{0}(\boldsymbol{x}, \boldsymbol{x}_{l})|^{2}\right)^{1/2}} \left(\hat{G}_{0}(\boldsymbol{x}, \boldsymbol{x}_{j})\right)_{j=1,...,n}$$

The matrix A_0 has rank 1 and its unique non-zero singular value is σ_{ref} .

Remark: other possible models: perfectly conducting crack, small inclusion, ...

Question 1: What is the structure of the (SVD of the) measured matrix **A** in the presence of noise ? When do we sound the alarm (presence of a reflector) ? CEMRACS August 3, 2011

Imaging functionals (1/3)

Let A be the (symmetrized) measured response matrix. Perform a SVD:

$$\mathbf{A} = \sum_{l=1}^n \sigma^{(l)} oldsymbol{v}^{(l)} (oldsymbol{v}^{(l)})^T$$

Remember that the unperturbed matrix is $\mathbf{A}_0 = \sigma_{\mathrm{ref}} \boldsymbol{g}(\boldsymbol{x}_{\mathrm{ref}}) \boldsymbol{g}(\boldsymbol{x}_{\mathrm{ref}})^T$.

• MUSIC functional (\mathbf{P} =projection on image space of \mathbf{A} =projection on span($\boldsymbol{v}^{(1)}$)):

$$\mathcal{I}_{ ext{MUSIC}}(oldsymbol{x}) = \left\| (\mathbf{I} - \mathbf{P}) oldsymbol{g}(oldsymbol{x})
ight\|^{-1} = \left\| oldsymbol{g}(oldsymbol{x}) - igl(oldsymbol{v}^{(1)}^T oldsymbol{g}(oldsymbol{x}) igr) oldsymbol{v}^{(1)}
ight\|^{-1}$$

• Reverse-Time migration functional:

$$\mathcal{I}_{\mathrm{RT}}(\boldsymbol{x}) = \overline{\boldsymbol{g}(\boldsymbol{x})}^T \mathbf{A} \overline{\boldsymbol{g}(\boldsymbol{x})}$$

• Kirchhoff Migration functional:

$$\mathcal{I}_{\mathrm{KM}}(\boldsymbol{x}) = \overline{\boldsymbol{d}(\boldsymbol{x})}^T \mathbf{A} \overline{\boldsymbol{d}(\boldsymbol{x})}$$

where

$$\boldsymbol{d}(\boldsymbol{x}) = \frac{1}{\sqrt{n}} \Big(\exp(i\frac{\omega}{c_0}|\boldsymbol{x} - \boldsymbol{x}_j|) \Big)_{j=1,\dots,n}$$

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Imaging functionals (2/3)

A general imaging functional can be obtained using a weighted-subspace migration:

$$egin{aligned} \mathcal{I}_{ ext{SM}}(oldsymbol{x},oldsymbol{w}) &= & \overline{oldsymbol{g}}(oldsymbol{x})^T \Big[\sum_{l=1}^n w_l(oldsymbol{x}) oldsymbol{v}^{(l)}(oldsymbol{v}^{(l)})^T \Big] \overline{oldsymbol{g}}(oldsymbol{x}) \ &= & \sum_{l=1}^n w_l(oldsymbol{x}) igg(oldsymbol{g}(oldsymbol{x})^T oldsymbol{v}^{(l)})^2 \end{aligned}$$

where $\boldsymbol{w}(\boldsymbol{x}) = (w_l(\boldsymbol{x}))_{l=1,...,n}$ are (complex) weights.

• Denote $w_l^{(1)}(\boldsymbol{x}) = \sigma^{(l)}$. Then $\mathcal{I}_{\mathrm{SM}}(\boldsymbol{x}, \boldsymbol{w}^{(1)})$ corresponds to Reverse-Time migration: $\mathcal{I}_{\mathrm{SM}}(\boldsymbol{x}, \boldsymbol{w}^{(1)}) = \mathcal{I}_{\mathrm{RT}}(\boldsymbol{x})$

• Denote $w_l^{(2)}(\boldsymbol{x}) = \exp\left(-i2 \arg\left(\overline{\boldsymbol{g}(\boldsymbol{x})}^T \boldsymbol{v}^{(1)}\right)\right) \mathbf{1}_1(l)$. Then $\mathcal{I}_{\mathrm{SM}}(\boldsymbol{x}, \boldsymbol{w}^{(2)})$ corresponds to MUSIC:

$$\begin{split} \mathcal{I}_{\text{MUSIC}}(\bm{x}) &= \|\bm{g}(\bm{x}) - \big(\overline{\bm{v}^{(1)}}^T \bm{g}(\bm{x})\big) \bm{v}^{(1)} \|^{-1} = \big(1 - \big|\overline{\bm{g}(\bm{x})}^T \bm{v}^{(1)}\big|^2\big)^{-1/2} \\ &= \big(1 - \mathcal{I}_{\text{SM}}(\bm{x}, \bm{w}^{(2)})\big)^{-1/2} \end{split}$$

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Imaging functionals (3/3)



Imaging functionals for a point reflector in the absence of noise. Left: Reverse-Time migration, center: Kirchhoff Migration, right: MUSIC (or more exactly, $1 - \mathcal{I}_{\text{MUSIC}}(\boldsymbol{x})^{-2}$).

Question 2: What is the best imaging functional to localize a reflector in the presence of noise ?

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A point reflector in a noisy environment: the response matrix

• In the presence of a reflector and noise, the response matrix **A** has the form

$$\mathbf{A} = \mathbf{A}_0 + \mathbf{W}$$

- The matrix A_0 is the rank-one matrix that corresponds to the reflector.
- The matrix **W** models noise.

Assume that:

there is additive measurement noise (and symmetrize the matrix, since the unperturbed response matrix is symmetric).

Then:

The matrix **W** is complex symmetric Gaussian, i.e., $W_{jl} = W_{lj}$ and W_{jl} , $j \leq l$ obey independent complex Gaussian random variables with mean zero and variance δ^2 off the diagonal $(j \neq l)$ and $2\delta^2$ on the diagonal (j = l).

The response matrix without reflector and with noise

• Denote

- $\sigma_1^{(n)} \ge \sigma_2^{(n)} \ge \sigma_3^{(n)} \ge \cdots \ge \sigma_n^{(n)}$ the singular values of the response matrix **A**. - $\sigma_{\text{noise}} = \sqrt{n\delta}$ where δ^2 is the variance of the entries of the matrix.

In the regime $n \gg 1$ (number of sensors $\gg 1$):

a) The singular values $(\sigma_j^{(n)})_{j=1,...,n}$ follow a quarter-circle distribution:

$$\frac{1}{n} \operatorname{Card}(j = 1, \dots, n, \sigma_{j}^{(n)} \in [a, b]) \xrightarrow{n \to \infty} \frac{1}{\sigma_{\operatorname{noise}}} \int_{a}^{b} \rho_{\operatorname{qc}}(\frac{\sigma}{\sigma_{\operatorname{noise}}}) d\sigma$$

$$\rho_{\operatorname{qc}}(\sigma) = \begin{cases} \frac{1}{\pi} \sqrt{4 - \sigma^{2}} & \text{if } 0 < \sigma \leq 2, \\ 0 & \text{otherwise.} \end{cases}$$



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In the regime $n \gg 1$ (number of sensors $\gg 1$):

b) The largest singular value satisfies

$$\sigma_1^{(n)} = \sigma_{\text{noise}} \left[2 + 2^{-2/3} n^{-2/3} Z_1 + o(n^{-2/3}) \right]$$

where Z_1 follows a type 1 Tracy-Widom distribution ($\mathbb{E}[Z_1] \simeq -1.21$, Var $(Z_1) \simeq 1.61$).



Histogram of $2^{2/3}n^{2/3}(\sigma_1^{(n)}/\sigma_{\text{noise}}-2)$ obtained from MC simulations with n = 50 in the absence of reflector (solid) and compared with the theoretical type 1 Tracy-Widom distribution (dashed).

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The response matrix with a point reflector and with noise

• Denote

- $\sigma_1^{(n)} \ge \sigma_2^{(n)} \ge \sigma_3^{(n)} \ge \cdots \ge \sigma_n^{(n)}$ the singular values of the response matrix **A**. - $\sigma_{\text{noise}} = \sqrt{n\delta}$ where δ^2 is the variance of the entries of the matrix.

In the regime $n \gg 1$ (number of sensors $\gg 1$):

a) The second singular value satisfies $\sigma_2^{(n)} \simeq 2\sigma_{\text{noise}}[1 + O(n^{-2/3})]$, the singular values $(\sigma_j^{(n)})_{j=2,...,n}$ follow a quarter-circle distribution.

The response matrix with a point reflector and with noise

• Denote

- $\sigma_1^{(n)} \ge \sigma_2^{(n)} \ge \sigma_3^{(n)} \ge \cdots \ge \sigma_n^{(n)}$ the singular values of the response matrix **A**. - $\sigma_{\text{noise}} = \sqrt{n\delta}$ where δ^2 is the variance of the entries of the matrix.

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a) The second singular value satisfies $\sigma_2^{(n)} \simeq 2\sigma_{\text{noise}}[1 + O(n^{-2/3})]$, the singular values $(\sigma_j^{(n)})_{j=2,...,n}$ follow a quarter-circle distribution.

b1) If $\sigma_{\text{noise}} < \sigma_{\text{ref}}$, then the largest singular value satisfies

$$\sigma_1^{(n)} = \sigma_{\rm ref} \left[1 + \sigma_{\rm noise}^2 \sigma_{\rm ref}^{-2} + n^{-1/2} \sigma_{\rm noise} \sigma_{\rm ref}^{-1} \left(1 - \sigma_{\rm noise}^2 \sigma_{\rm ref}^{-2} \right)^{1/2} Z_0 + o(n^{1/2}) \right]$$

where Z_0 follows a standard Gaussian distribution.

b2) If $\sigma_{\text{noise}} > \sigma_{\text{ref}}$, then the largest singular value satisfies

$$\sigma_1^{(n)} = \sigma_{\text{noise}} \left[2 + 2^{-2/3} n^{-2/3} Z_1 + o(n^{-2/3}) \right]$$

where Z_1 follows a type 1 Tracy-Widom distribution ($\mathbb{E}[Z_1] \simeq -1.21$, Var $(Z_1) \simeq 1.61$).

Proof: random matrix theory.



Mean first and second singular values.

Consider

$$R := \frac{\sigma_1^{(n)}}{\left(\frac{1}{n-2}\sum_{j=2}^n (\sigma_j^{(n)})^2\right)^{1/2}}$$

In the regime $n \gg 1$ (number of sensors $\gg 1$)

• In the absence of reflector,

$$R \stackrel{dist.}{=} 2 + \frac{1}{2^{2/3} n^{2/3}} Z_1$$

where Z_1 follows a type 1 Tracy-Widom distribution.

• In the presence of a reflector, if $\sigma_{\rm ref} > \sigma_{\rm noise}$, then

$$R \stackrel{dist.}{=} \frac{\sigma_{\text{ref}}}{\sigma_{\text{noise}}} + \frac{\sigma_{\text{noise}}}{\sigma_{\text{ref}}} + \frac{1}{\sqrt{n}}\sqrt{1 - \sigma_{\text{noise}}^2 \sigma_{\text{ref}}^{-2}} Z_0$$

where Z_0 follows a standard Gaussian distribution.

If $\sigma_{\text{ref}} < \sigma_{\text{noise}}$ then $R \stackrel{dist.}{=} 2 + \frac{1}{2^{2/3}n^{2/3}}Z_1$, as in the absence of a reflector !

• Detection test with level r: If the data gives R, then sound the alarm if R > r.

• The false alarm rate (FAR) is the probability to sound the alarm when there is no reflector:

 $FAR = \mathbb{P}(R > r)$ without reflector)

The probability of detection (POD) is the probability to sound the alarm when there is a reflector:

 $POD = \mathbb{P}(R > r | \text{ with reflector })$

• Detection test with level r: If the data gives R, then sound the alarm if R > r.

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• Fix $\alpha \in (0, 1)$. Choose

$$r_{\alpha} = 2 + \frac{1}{2^{2/3}n^{2/3}}\Phi_{\mathrm{TW1}}^{-1}(1-\alpha),$$

where Φ_{TW1} is the type 1 Tracy-Widom cumulative distribution function. For instance, $\Phi_{\text{TW1}}^{-1}(0.95) \simeq 0.98$. Main results:

- The FAR of the test $R > r_{\alpha}$ is α .
- The POD of the test $R > r_{\alpha}$ is

$$\text{POD} = \max\left\{\alpha, \Phi\left(\sqrt{n}\frac{\frac{\sigma_{\text{ref}}}{\sigma_{\text{noise}}} + \frac{\sigma_{\text{noise}}}{\sigma_{\text{ref}}} - r_{\alpha}}{\sqrt{1 - (\sigma_{\text{noise}}/\sigma_{\text{ref}})^2}}\right)\right\}$$

where Φ is the standard Gaussian cumulative distribution function.

• By Neyman-Pearson lemma the test $R > r_{\alpha}$ maximizes the POD for a given FAR α . CEMRACS August 3, 2011 The POD increases with the number n of sensors, with $\sigma_{\rm ref}/\sigma_{\rm noise}$, and with the FAR α .



The SV-based test becomes powerful when $\sigma_{\rm ref} > \sigma_{\rm noise}$.

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Migration-based detection test

Reverse-Time imaging functional:

$$\mathcal{I}_{\mathrm{RT}}(\boldsymbol{x}) = \overline{\boldsymbol{g}(\boldsymbol{x})}^T \mathbf{A} \overline{\boldsymbol{g}(\boldsymbol{x})} \qquad \text{with } \boldsymbol{g}(\boldsymbol{x}) = \frac{1}{\left(\sum_{l=1}^n |\hat{G}_0(\boldsymbol{x}, \boldsymbol{x}_l)|^2\right)^{1/2}} \left(\hat{G}_0(\boldsymbol{x}, \boldsymbol{x}_j)\right)_{j=1,\dots,n}$$

• Imaging of a point reflector without noise: $\mathcal{I}_{\mathrm{RT}}$ is a peak centered at x_{ref} :

$$\mathcal{I}_{\mathrm{RT}}(\boldsymbol{x}) = \sigma_{\mathrm{ref}} h(\boldsymbol{x} - \boldsymbol{x}_{\mathrm{ref}}) = \sigma_{\mathrm{ref}} \left(\overline{\boldsymbol{g}(\boldsymbol{x})}^T \boldsymbol{g}(\boldsymbol{x}_{\mathrm{ref}})\right)^2$$

 $h = \text{point spread function}; \text{ maximal at } \mathbf{0} \text{ (Cauchy-Schwartz)}; h(\mathbf{0}) = 1.$

• Full aperture:

$$h(\boldsymbol{x}) = \operatorname{sinc}^2\left(\frac{\pi|\boldsymbol{x}|}{\lambda_0}\right)$$

 \hookrightarrow the width of the function $h(\boldsymbol{x})$ is $\lambda_0/2$ (diffraction limit).

• Imaging of noise without reflector: \mathcal{I}_{RT} is a speckle pattern, i.e. a stationary Gaussian random field with mean zero, variance $2\delta^2$, and covariance function:

$$\mathbb{E}\big[\mathcal{I}_{\mathrm{RT}}(\boldsymbol{x})\overline{\mathcal{I}_{\mathrm{RT}}}(\boldsymbol{y})\big] = 2\delta^2 h(\boldsymbol{x}-\boldsymbol{y})$$

The hotspot volume is defined by

$$V_c = \frac{\pi^{3/2}}{(\det \mathbf{H})^{1/2}}, \qquad \mathbf{H} = \left(-\partial_{x_j x_l}^2 h(\mathbf{0})\right)_{j,l=1,...,3}.$$

• Full aperture:

$$\mathbf{H} = \frac{8\pi^2}{3\lambda_0^2} \mathbf{I}, \qquad V_c = \frac{3^{3/2}}{(2\pi)^{3/2}} \lambda_0^3$$





With reflector, without noise Point spread function

Without reflector, with noise Speckle pattern



With reflector, with noise

• Consider

$$S := \frac{\|\mathcal{I}_{\mathrm{RT}}\|_{L^{\infty}(\Omega)}^2 |\Omega|}{\|\mathcal{I}_{\mathrm{RT}}\|_{L^2(\Omega)}^2}.$$

In the regime $n \gg 1$ (number of sensors $\gg 1$) and $|\Omega| \gg V_c$ (search volume \gg hotspot volume):

a) In the absence of a reflector, then

$$S \stackrel{dist.}{=} \ln \frac{|\Omega|}{V_c} + \frac{3}{2} \ln \ln \frac{|\Omega|}{V_c} + Z,$$

where Z follows a Gumbel distribution.

b) In the presence of a reflector, then

$$S \stackrel{dist.}{=} \max\Big\{\frac{\sigma_{\text{ref}}^2}{2\delta^2} + \frac{\sigma_{\text{ref}}}{\delta}Z_0, \ln\frac{|\Omega|}{V_c} + \frac{3}{2}\ln\ln\frac{|\Omega|}{V_c} + Z\Big\},\$$

where Z_0 follows a standard Gaussian distribution and is independent of Z.

Proof: extreme value theory for random fields.

• Migration-based detection test: If the data gives S, sound the alarm if S > s. Fix $\alpha \in (0, 1)$. Choose

$$s_{\alpha} = \ln \frac{|\Omega|}{V_c} + \frac{3}{2} \ln \ln \frac{|\Omega|}{V_c} + \Phi_{\rm G}^{-1}(1-\alpha),$$

where $\Phi_{\rm G}(x) = \exp(-e^{-x})$ is the Gumbel cumulative distribution function. For instance $\Phi_{\rm G}^{-1}(0.95) \simeq 2.97$.

Main results:

- The false alarm rate (FAR) of the test $S > s_{\alpha}$ is α .
- The probability of detection (POD) of the test $S > s_{\alpha}$ is

$$\text{POD} = \max\left\{\Phi\left(\frac{\frac{\sigma_{\text{ref}}^2}{2\delta^2} - s_{\alpha}}{\frac{\sigma_{\text{ref}}}{\delta}}\right), \alpha\right\}$$

where Φ is the standard Gaussian cumulative distribution function.

• By Neyman Pearson lemma, the test $S > s_{\alpha}$ maximizes the POD for a given FAR α .



The migration-based test becomes powerful when $\sigma_{\rm ref} > \sqrt{2} \delta \ln^{1/2} (|\Omega|/V_c)$.

Remember: $\sigma_{\text{noise}} = \sqrt{n}\delta$.

Comparison between SV-based test and migration-based test

- The migration-based test becomes powerful when $\sigma_{\rm ref} > \sqrt{2} \delta \ln^{1/2} (|\Omega|/V_c)$.
- The SV-based test becomes powerful when $\sigma_{\rm ref} > \delta \sqrt{n}$.
- Therefore the migration-based test is more (resp. less) powerful than the SV-based test when $n > (\text{resp.} <) 2 \ln(|\Omega|/V_c)$.
- In practice, we usually have $n > 2 \ln(|\Omega|/V_c)$, and therefore the migration-based test is more efficient than the SV-based test.
- The SV-based test is simpler to implement than the migration-based test.

Optimal migration for localization

• Given the parameters \boldsymbol{x} (the reflector location), σ (the reflector singular value), and δ^2 (the noise variance), the likelihood of the observations \mathbf{A} is proportional to

$$l_0(\mathbf{A} \mid \boldsymbol{x}, \sigma, \delta^2) = \frac{1}{\delta^{n^2 + n}} \exp\left(-\frac{\left\|\mathbf{A} - \sigma \boldsymbol{g}(\boldsymbol{x})\boldsymbol{g}(\boldsymbol{x})^T\right\|_F^2}{2\delta^2}\right)$$

• Prior information on $(\boldsymbol{x}, \sigma, \delta^2)$: nothing (Jeffrey prior).

• Bayes theorem: Given the observations \mathbf{A} , the likelihood function of the parameters \boldsymbol{x} (the reflector location), σ (the reflector singular value), and δ^2 (the noise variance) is proportional to

$$l_0(\boldsymbol{x}, \sigma, \delta^2 \mid \mathbf{A}) = \frac{1}{\delta^{n^2 + n + 1}} \exp\left(-\frac{\left\|\mathbf{A} - \sigma \boldsymbol{g}(\boldsymbol{x})\boldsymbol{g}(\boldsymbol{x})^T\right\|_F^2}{2\delta^2}\right)$$

• The maximum likelihood estimate \hat{x} of x (and the nuisance parameters δ^2 and σ) is found by maximizing the likelihood function. Main result:

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\hat{oldsymbol{x}} = rgmax_{oldsymbol{x}} |\mathcal{I}_{	ext{RT}}(oldsymbol{x})|^2
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 \hookrightarrow Reverse-Time migration is the best method in the presence of additive noise.

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Proof. Given the observations **A**, the likelihood function of the parameters \boldsymbol{x} (the reflector location), σ (the reflector singular value), and δ^2 (the noise variance) is proportional to

$$l_0(\boldsymbol{x}, \sigma, \delta^2 \mid \mathbf{A}) = \frac{1}{\delta^{n^2 + n + 1}} \exp\left(-\frac{\left\|\mathbf{A} - \sigma \boldsymbol{g}(\boldsymbol{x}) \boldsymbol{g}(\boldsymbol{x})^T\right\|_F^2}{2\delta^2}\right)$$

The maximum likelihood estimate \hat{x} of x (and the nuisance parameters δ^2 and σ) is found by maximizing the likelihood function:

$$ig(\hat{oldsymbol{x}},\hat{\sigma},\hat{\delta}^2ig) = rgmax_{oldsymbol{x},\sigma,\delta^2} l_0ig(oldsymbol{x},\sigma,\delta^2\mid \mathbf{A}ig).$$

We first eliminate δ^2 by requiring

$$\frac{\partial l_0(\boldsymbol{x},\sigma,\delta^2 \mid \mathbf{A})}{\partial \delta} = 0.$$

This gives

$$\hat{\delta}^2 = \frac{\|\mathbf{A} - \sigma \boldsymbol{g}(\boldsymbol{x}) \boldsymbol{g}(\boldsymbol{x})^T\|_F^2}{n^2 + n + 1},$$

and the likelihood ratio is then proportional to

$$l_0(\boldsymbol{x},\sigma,\hat{\delta}^2 \mid \mathbf{A}) \simeq \left\| \mathbf{A} - \sigma \boldsymbol{g}(\boldsymbol{x}) \boldsymbol{g}(\boldsymbol{x})^T \right\|_F^{-(n^2+n+1)/2}.$$

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Since **A** is complex symmetric it admits a symmetric SVD: there exist unitary vectors $\boldsymbol{v}^{(l)}$ and nonnegative numbers $\sigma^{(l)}$ (the singular values) such that

$$\mathbf{A} = \sum_{l=1}^n \sigma^{(l)} oldsymbol{v}^{(l)} oldsymbol{v}^{(l)}^T$$

We can write

$$\begin{aligned} \left\| \mathbf{A} - \sigma \boldsymbol{g}(\boldsymbol{x}) \boldsymbol{g}(\boldsymbol{x})^T \right\|_F^2 &= \left\| \tilde{\boldsymbol{v}} - \sigma \tilde{\boldsymbol{g}}(\boldsymbol{x}) \right\|_2^2 \\ \text{for } \tilde{\boldsymbol{v}} &= \sum_{l=1}^n \sigma^{(l)} \boldsymbol{v}^{(l)} \otimes \boldsymbol{v}^{(l)} \text{ and } \tilde{\boldsymbol{g}}(\boldsymbol{x}) = \boldsymbol{g}(\boldsymbol{x}) \otimes \boldsymbol{g}(\boldsymbol{x}). \text{ Using } \| \tilde{\boldsymbol{g}}(\boldsymbol{x}) \|_2 = \| \boldsymbol{g}(\boldsymbol{x}) \|^2 = 1, \\ \hat{\sigma} &= \operatorname*{argmin}_{\sigma} \| \tilde{\boldsymbol{v}} - \sigma \tilde{\boldsymbol{g}}(\boldsymbol{x}) \|_2^2 = \overline{\tilde{\boldsymbol{g}}(\boldsymbol{x})}^T \tilde{\boldsymbol{v}} \end{aligned}$$

Therefore the estimate \hat{x} derives from maximizing the MUSIC-type function

$$\hat{oldsymbol{x}} = rgmin_{oldsymbol{x}} \left\| ilde{oldsymbol{v}} - ig(\overline{oldsymbol{g}}(oldsymbol{x})^T ilde{oldsymbol{v}} ig) ilde{oldsymbol{g}}(oldsymbol{x})
ight\|_2^2$$

Note however that \hat{x} is not the maximizer of the MUSIC functional since all singular vectors (weighted by the singular values) contribute to \tilde{v} . We have in fact

$$\begin{split} \left\| \tilde{\boldsymbol{v}} - \left(\overline{\tilde{\boldsymbol{g}}(\boldsymbol{x})}^T \tilde{\boldsymbol{v}} \right) \tilde{\boldsymbol{g}}(\boldsymbol{x}) \right\|_2^2 &= \left\| \tilde{\boldsymbol{v}} \right\|_2^2 - \left| \overline{\tilde{\boldsymbol{g}}(\boldsymbol{x})}^T \tilde{\boldsymbol{v}} \right|^2 \quad = \quad \left\| \tilde{\boldsymbol{v}} \right\|_2^2 - \left| \sum_{l=1}^n \sigma^{(l)} \left(\overline{\boldsymbol{g}(\boldsymbol{x})}^T \boldsymbol{v}^{(l)} \right)^2 \right|^2 \\ &= \quad \left\| \mathbf{A} \right\|_F^2 - \left| \mathcal{I}_{\mathrm{RT}}(\boldsymbol{x}) \right|^2 \end{split}$$

This gives $\hat{\boldsymbol{x}} = \operatorname*{argmin}_{\boldsymbol{x}} (\| \mathbf{A} \|_F^2 - |\mathcal{I}_{\mathrm{RT}}(\boldsymbol{x})|^2).$

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Optimal migration for localization

Important remark:

The Bayesian localization scheme can be used once the detection test has passed. Bayesian analysis is powerful but depends on the prior.

Here the prior is: there exists a reflector.

Statistical analysis of localization error

Optimal estimator:

$$\hat{oldsymbol{x}} = rgmax_{oldsymbol{x}} \left| \mathcal{I}_{ ext{RT}}(oldsymbol{x})
ight|^2$$

To leading order in $\delta/\sigma_{\rm ref}$, the estimator \hat{x} is unbiased and its covariance matrix is

$$\mathbb{E}ig[(\hat{oldsymbol{x}}-oldsymbol{x}_{ ext{ref}})(\hat{oldsymbol{x}}-oldsymbol{x}_{ ext{ref}})^Tig]=rac{\delta^2}{\sigma_{ ext{ref}}^2}\mathbf{H}^{-1}.$$

Full aperture: $\mathbb{E}\left[(\hat{x}_j - x_{\mathrm{ref},j})^2\right] = \frac{\delta^2}{\sigma_{\mathrm{ref}}^2} \frac{3}{2\pi^2} \lambda_0^2, \quad j = 1, \dots, 3.$

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Imaging functionals in the absence of noise.

Left: reverse-time migration, center: Kirchhoff migration, right: MUSIC.



Imaging functionals in the presence of noise.

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Standard deviation of the estimated reflector location obtained with three different imaging methods (here $\lambda_0 = 1$, n = 100).

Conclusions

• Statistical and stochastic tools:

- Random matrix theory: description of the distribution of the singular values of the array response matrix in a noisy environment.

- Extreme value theory: description of the speckle pattern obtained by migration (backpropagation) of the array data.

- Bayesian analysis: optimal localization of a target.

• Optimal tests involve non-Gaussian distributions (Gumbel, Tracy-Widom).

• What is important is the structure of the response matrix (symmetric, Hermitian, Hankel, Toeplitz, ...), not much the marginal distribution of the entries.

- It is possible to extend the results to
- several reflectors, cracks, or inclusions,

- other noisy environments (random medium in the single-scattering regime for instance).

The main hypothesis is that the information is low-rank while the noise is high-rank.

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