

Linear Recovery from Gaussian Observations

joint work with

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http://www2.isye.gatech.edu/~nemirovs/StatOpt_LN.pdf

Paris, January 9, 2017

Situation: “In the nature” there exists a signal x known to belong to a given convex compact set $\mathcal{X} \subset \mathbb{R}^n$. We observe corrupted by noise affine image of the signal:

$$\omega = Ax + \sigma\xi \in \Omega = \mathbb{R}^m$$

- A : given $m \times n$ sensing matrix
- ξ : random observation noise
- **Our goal** is to recover the image Bx of x under a given affine mapping $B: \mathbb{R}^n \rightarrow \mathbb{R}^k$.
- **Risk** of a candidate estimate $\hat{x}(\cdot) : \Omega \rightarrow \mathbb{R}^k$ is defined as

$$\text{Risk}[\hat{x}|\mathcal{X}] = \sup_{x \in \mathcal{X}} \sqrt{\mathbf{E}_{\xi} \{ \|Bx - \hat{x}(Ax + \sigma\xi)\|_2^2 \}}$$

\Rightarrow Risk² is the worst-case, over $x \in \mathcal{X}$, expected $\|\cdot\|_2^2$ recovery error.

Agenda: Under appropriate assumptions on \mathcal{X} , we are to show that

- *One can build, in a computationally efficient fashion, the (nearly) best, in terms of risk, estimate from the family of linear estimates*

$$\hat{x}(\omega) = \hat{x}_H(\omega) = H^T \omega \quad [H \in \mathbb{R}^{m \times k}]$$

- *The resulting linear estimate is nearly optimal among **all** estimates, linear and nonlinear alike.*

Linear estimation of signal in Gaussian noise

- ...
- Kuks & Olman, 1971, 1972
- Rao 1972, 1973, Pilz, 1981, 1986, ..., Drygas, 1996, Christopeit & Helmes, 1996, Arnold & Stahlecker, 2000, ...
- Pinsker 1980, Efromovich & Pinsker, 1981, 1982, Efromovich & Pinsker 1996, Golubev Levit & Tsybakov, 1996, ..., Efromovich, 2008, ...
- Donoho, Liu, McGibbon, 1990
- ...

Risk of linear estimation

Assuming that ξ is zero mean with unit covariance matrix, we can easily compute the risk of a linear estimate $\hat{x}_H(\omega) = H^T \omega$

$$\begin{aligned}\text{Risk}^2[\hat{x}_H|\mathcal{X}] &= \max_{x \in \mathcal{X}} \mathbf{E}_\xi \left\{ \| [B - H^T A]x - \sigma H^T \xi \|_2^2 \right\} \\ &= \max_{x \in \mathcal{X}} \left\{ \| [B - H^T A]x \|_2^2 + \sigma^2 \mathbf{E}_\xi \{ \text{Tr}(H^T \xi \xi^T H) \} \right\} \\ &= \sigma^2 \text{Tr}(H^T H) + \max_{x \in \mathcal{X}} \text{Tr}(xx^T [B^T - A^T H][B - H^T A]).\end{aligned}$$

Note: ϕ is convex \Rightarrow building the minimum risk linear estimate reduces to solving convex minimization problem

$$\text{Opt}^P = \min_H \left[\phi(H) = \max_{x \in \mathcal{X}} \text{Tr}(xx^T [B^T - A^T H][B - H^T A]) + \sigma^2 \text{Tr}(H^T H) \right]. \quad (*)$$

Convex function ϕ is given implicitly and can be difficult to compute, making (*) difficult as well.

Fact: essentially, the only cases when (*) is known to be easy are those when

- \mathcal{X} is given as a convex hull of finite set of moderate cardinality
- \mathcal{X} is an ellipsoid: for $W \in \mathbf{S}^n$ and $S \succ 0$

$$\max_{x^T S x \leq 1} \text{Tr}(x x^T W) = \lambda_{\max}(S^{-1}W).$$

where $\lambda_{\max}(\cdot)$ is the maximal eigenvalue.

When \mathcal{X} is a “box,” computing Ψ is NP-hard...

- When Ψ is difficult to compute, we can to replace Ψ in the design problem (*) with an efficiently computable convex upper bound $\Psi^+(H)$.
- We are about to consider a family of sets \mathcal{X} – *ellitopes* – for which reasonably tight bounds Ψ^+ of desired type are available.

An ellitope is a set $\mathcal{X} \subset \mathbb{R}^n$ given as

$$\mathcal{X} = \{x \in \mathbb{R}^n : \exists y \in \mathbb{R}^N, t \in \mathcal{T} : x = Py, y^T S_k y \leq t_k, 1 \leq k \leq K\}$$

where

- P is a given $n \times N$ matrix (we can assume that $P = I_n$),
- $S_k \succeq 0$ are positive semidefinite matrices with $\sum_k S_k \succ 0$
- \mathcal{T} is a convex compact subset of K -dimensional nonnegative orthant \mathbb{R}_+^K such that
 - \mathcal{T} contains some positive vectors
 - \mathcal{T} is *monotone*: if $0 \leq t' \leq t$ and $t \in \mathcal{T}$, then $t' \in \mathcal{T}$ as well.

Note: every *ellitope* is a symmetric w.r.t. the origin convex compact set.

Examples

[A.] A centered at the origin ellipsoid ($K = 1$, $\mathcal{T} = [0; 1]$)

[B.] (Bounded) intersection of K ellipsoids/elliptic cylinders centered at the origin
($\mathcal{T} = \{t \in \mathbb{R}^K : 0 \leq t_k \leq 1, k \leq N\}$)

[C.] Box $\{x \in \mathbb{R}^n : -1 \leq x_i \leq 1\}$ ($\mathcal{T} = \{t \in \mathbb{R}^n : 0 \leq t_k \leq 1, k \leq K = n\}$, $x^T S_k x = x_k^2$)

[D.] $\mathcal{X} = \{x \in \mathbb{R}^n : \|x\|_p \leq 1\}$ with $p \geq 2$

($\mathcal{T} = \{t \in \mathbb{R}_+^n : \|t\|_{p/2} \leq 1\}$, $x^T S_k x = x_k^2$, $k \leq K = n$)

Ellitopes admit fully algorithmic calculus: if \mathcal{X}_i , $1 \leq i \leq I$, are ellitopes, so are

- linear images of \mathcal{X}_i
- $\mathcal{X}_1 \times \dots \times \mathcal{X}_I$
- inverse linear images of \mathcal{X}_i under linear embeddings
- $\text{Conv}(\bigcup_i \mathcal{X}_i)$
- $\mathcal{X}_1 + \dots + \mathcal{X}_I$
- $\bigcap_i \mathcal{X}_i$
- ...

Observation

Let

$$\mathcal{X} = \{x \in \mathbb{R}^n : \exists t \in \mathcal{T} : x^T S_k x \leq t_k, 1 \leq k \leq K\}$$

be an ellitope. Given a quadratic form $x^T W x$, $W \in \mathbf{S}^n$, one has

$$\max_{x \in \mathcal{X}} x^T W x = \max_{x \in \mathcal{X}} \text{Tr}(x x^T W) \leq \max_{Q \in \mathcal{Q}} \text{Tr}(Q W),$$

where

$$\mathcal{Q} := \{Q \in \mathbf{S}^n : Q \succeq 0, \exists t \in \mathcal{T} : \text{Tr}(Q S_k) \leq t_k, k \leq K\}.$$

We conclude that

$$\phi(H) \leq \varphi(H) := \sigma^2 \text{Tr}(H^T H) + \max_{Q \in \mathcal{Q}} \text{Tr}(Q(A^T H - B^T)(H^T A - B)),$$

and

$$\text{Risk}^2[\hat{x}_H | \mathcal{X}] \leq \min_H \varphi(H).$$

This attracts our attention to the optimization problem

$$\text{Opt}^P = \min_H \left\{ \varphi(H) = \max_{Q \in \mathcal{Q}} \underbrace{\left[\sigma^2 \text{Tr}(H^T H) + \text{Tr}(Q(A^T H - B^T)(H^T A - B)) \right]}_{\Phi(H, Q)} \right\}. \quad (P)$$

Note that (P) is the primal problem

$$\min_H \left[\max_{Q \in \mathcal{Q}} \Phi(H, Q) \right]$$

associated with the convex-concave saddle point function $\Phi(H, Q)$. The **dual problem** associated with $\Phi(H, Q)$ is

$$\max_{Q \in \mathcal{Q}} \left[\min_H \Phi(H, Q) \right],$$

that is, the problem

$$\text{Opt}^D = \max_{Q \in \mathcal{Q}} \left\{ \psi(Q) := \min_H \left[\sigma^2 \text{Tr}(H^T H) + \text{Tr}(Q(A^T H - B^T)(H^T A - B)) \right] \right\}. \quad (D)$$

By the Sion-Kakutani theorem, (P) and (D) are solvable with equal optimal values: $\text{Opt}^D = \text{Opt}^P = \text{Opt}$.

Note that the minimizer of $\Phi(\cdot, Q)$ can be easily computed:

$$H(Q) = (\sigma^2 I_m + AQA^T)^{-1} AQB^T,$$

so that

$$\psi(Q) = \text{Tr}(B[Q - QA^T(\sigma^2 I_m + AQA^T)^{-1}AQ]B^T),$$

and the dual problem reads

$$\text{Opt} = \max_{Q,t} \left\{ \text{Tr}(B[Q - QA^T(\sigma^2 I_m + AQA^T)^{-1}AQ]B^T), \right. \\ \left. Q \succeq 0, t \in \mathcal{T}, \text{Tr}(QS_k) \leq t_k, k \leq K \right\} \quad (D)$$

In fact, both (P) and (D) can be cast as **Semidefinite Optimization problems**.

In particular, (P) can be rewritten as

$$\text{Opt} = \min_{H,\lambda} \left\{ \sigma^2 \text{Tr}(H^T H) + \phi_{\mathcal{T}}(\lambda) : \begin{bmatrix} \sum_k \lambda_k S_k & B^T - A^T H \\ B - H^T A & I_\nu \end{bmatrix} \succeq 0, \lambda \geq 0 \right\} \quad (P)$$

where $\phi_{\mathcal{T}} : \mathbb{R}^K \rightarrow \mathbb{R}$ is the support function of \mathcal{T} :

$$\phi_{\mathcal{T}}(\lambda) = \max_{t \in \mathcal{T}} \lambda^T t.$$

Note that (P) is efficiently solvable whenever \mathcal{T} is computationally tractable.

Bottom line: Given matrices $A \in \mathbb{R}^{m \times n}$, $B \in \mathbb{R}^{k \times n}$ and an ellitope

$$\mathcal{X} = \{x \in \mathbb{R}^n : \exists t \in \mathcal{T} : x^T S_k x \leq t_k, 1 \leq k \leq K\} \quad (*)$$

consider the convex optimization problems

$$\text{Opt}^P = \min_H \varphi(H) \quad \text{and} \quad \text{Opt}^D = \max_{Q \in \mathcal{Q}} \psi(Q),$$

where $\mathcal{Q} := \{Q \in \mathbf{S}^n : Q \succeq 0, \exists t \in \mathcal{T} : \text{Tr}(Q S_k) \leq t_k, k \leq K\}$.

- The optimal values of two problems coincide, $\text{Opt}^P = \text{Opt}^D = \text{Opt}$.
- When noise ξ satisfies $\mathbf{E}\{\xi\} = 0$, and $\mathbf{E}\{\xi\xi^T\} = I_m$, the risk of the linear estimate $\hat{x}_{H_*}(\cdot)$ induced by the optimal solution H_* to the problem (this solution clearly exists provided $\sigma > 0$) satisfies the risk bound

$$\text{Risk}[\hat{x}_{H_*} | \mathcal{X}] \leq \sqrt{\text{Opt}}.$$

Bayesian risks

- *Minimax risk* $\text{Risk}[\hat{x}|\mathcal{X}]$ is defined as the *worst*, over the signals of interest, performance of $\hat{x}(\cdot)$
- *Bayesian risk* is the *average performance*, with the average taken over some *prior* probability distribution on the signals.

For the problem of $\|\cdot\|_2$ -recovering Bx via noisy observation

$$\omega = Ax + \sigma\xi, \quad \xi \sim P$$

this alternative reads as follows:

- (!) Given a probability distribution π of signal $x \in \mathbb{R}^n$, find an estimate $\hat{x}(\cdot)$ which minimizes

$$\text{Risk}^2(\hat{x}|\pi) := \int_{\pi} \left\{ \int_{\mathbb{R}^m} \|Bx - \hat{x}(Ax + \sigma\xi)\|_2^2 P(d\xi) \right\} \pi(dx)$$

- the average, over the distribution π of signals x , of expected $\|\cdot\|_2^2$ recovery error of Bx via observation $Ax + \sigma\xi$.

Let $P_{x,\omega}$ be the induced by π and P_ξ *joint distribution* of $(x, \omega = Ax + \sigma\xi)$ on $\mathbb{R}_x^n \times \mathbb{R}_\omega^m$. $P_{x,\omega}$ gives rise to

- *marginal distribution P_ω of ω ,*
- *conditional distribution $P_{x|\omega}$ of x given ω .*

We have

$$\begin{aligned} \text{Risk}^2(\hat{x}|\pi) &= \int_{\mathbb{R}_x^n \times \mathbb{R}_\omega^m} \|Bx - \hat{x}(\omega)\|_2^2 P_{x,\omega}(dx, d\omega) \\ &= \int_{\mathbb{R}^m} \left\{ \int_{\mathbb{R}^n} \|Bx - \hat{x}(\omega)\|_2^2 P_{x|\omega}(dx) \right\} P_\omega(d\omega) \end{aligned}$$

Assuming that the probability distribution π possesses finite second moments, one has

$$\min_{\hat{x}(\cdot)} \int_{\mathbb{R}^n \times \mathbb{R}^m} \|Bx - \hat{x}(\omega)\|_2^2 P_{x,\omega}(dx) = \int_{\mathbb{R}^n \times \mathbb{R}^m} \|Bx - \hat{x}_*(\omega)\|_2^2 P_{x,\omega}(dx),$$

where

$$\hat{x}_*(\omega) = \int_{\mathbb{R}^n} Bx P_{x|\omega}(dx).$$

Corollary [Gauss-Markov theorem]: Let $x \in \mathbb{R}^n$ and $\xi \in \mathbb{R}^m$ be independent zero-mean Gaussian random vectors. Assuming $\sigma > 0$ and the covariance matrix of ξ to be positive definite,

- conditional, given ω , distribution of x is normal, so that the conditional expectation $\hat{x}_*(\omega)$ is a linear function of ω ,
- as a result, an optimal solution $\hat{x}_*(\cdot)$ to the risk minimization problem

$$\min_{\hat{x}(\cdot)} \mathbf{E}_{x,\xi} \left\{ \|Bx - \hat{x}(Ax + \sigma\xi)\|_2^2 \right\}$$

exists and is a linear function of $\omega = Ax + \sigma\xi$.

In particular, when $\xi \sim \mathcal{N}(0, I_m)$ and $x \sim \mathcal{N}(0, Q)$, one has

$$\begin{aligned} \hat{x}_*(\omega) &= [\sigma^2 I_m + AQA^T]^{-1} AQB^T \omega \\ \text{Risk}^2(\hat{x}_* | \mathcal{N}(0, Q)) &= \text{Tr}(B[Q - QA^T[\sigma^2 I_m + AQA^T]^{-1}AQ]B^T) \end{aligned}$$

Course of actions (Pinsker's program)

- Let $\mathcal{N}(0, Q)$ be a Gaussian prior for the signal x which “sits on \mathcal{X} with high probability.” Then by the Gauss-Markov theorem the (“slightly reduced”) quantity

$$\varphi(Q) = \text{Tr}(B[Q - QA^T[\sigma^2 I_m + AQA^T]^{-1}AQ]B^T)$$

would be a lower bound on $\text{Risk}_{\text{Opt}}^2$.

- Note that $\mathbf{E}_{\eta \sim \mathcal{N}(0, Q)}\{\eta^T S \eta\} = \text{Tr}(SQ)$. Thus, selecting $Q \succeq 0$ according to

$$\exists t \in \mathcal{T} : \text{Tr}(QS_k) \leq t_k, k \leq K$$

we ensure that $\eta \sim \mathcal{N}(0, Q)$ sits in \mathcal{X} “on average.”

Imposing on $Q \succeq 0$ restriction

$$\exists t \in \mathcal{T} : \text{Tr}(QS_k) \leq \rho t_k, k \leq K, \quad [\rho > 0]$$

we enforce $\eta \sim \mathcal{N}(0, Q)$ to take values in \mathcal{X} with probability controlled by ρ and approaching 1 as $\rho \rightarrow +0$.

- The above considerations give rise to parametric optimization problem

$$\text{Opt}_*(\rho) = \max_{Q \succeq 0} \{\varphi(Q) : \exists t \in \mathcal{T} : \text{Tr}(QS_k) \leq \rho t_k, 1 \leq k \leq K\} \quad (P_\rho)$$

We *may expect* that for small ρ a “slightly corrected” $\text{Opt}_*(\rho)$ is a lower bound on $\text{Risk}_{\text{Opt}}^2$.

- As we have just seen, $\text{Opt}_*(1) = \text{Opt}$ (!). Since the optimal value of the (concave) optimization problem (P_ρ) is a concave function of ρ , we have

$$\text{Opt}_*(\rho) \geq \rho \text{Opt}, \quad 0 < \rho < 1.$$

Now, all we need is a simple result as follows:

Lemma Let S and Q be positive semidefinite $n \times n$ matrices with $\rho := \text{Tr}(SQ) \leq 1$, and let $\eta \sim \mathcal{N}(0, Q)$. Then

$$\text{Prob} \{ \eta^T S \eta > 1 \} \leq e^{-\frac{1-\rho+\rho \ln(\rho)}{2\rho}}$$

We arrive at the following

Theorem. Let us associate with ellitope $\mathcal{X} = \{x \in \mathbb{R}^n : \exists t \in \mathcal{T} : x^T S_k x \leq t_k, k \leq K\}$ the convex compact set

$$\mathcal{Q} = \{Q \in \mathbf{S}^n : Q \succeq 0, \exists t \in \mathcal{T} : \text{Tr}(QS_k) \leq t_k, k \leq K\},$$

and the quantity

$$M_* = \max_{Q \in \mathcal{Q}} \sqrt{\text{Tr}(BQB^T)}.$$

Then the linear estimate $\hat{x}_{H_*}(\omega) = H_*^T \omega$ of Bx , $x \in \mathcal{X}$, via observation $\omega = Ax + \sigma\xi$, $\xi \sim \mathcal{N}(0, I_m)$, given by the optimal solution H_* to the convex optimization problem

$$\text{Opt} = \min_{H, \lambda} \left\{ \phi_{\mathcal{T}}(\lambda) + \sigma^2 \text{Tr}(HH^T) : \begin{array}{l} \lambda \geq 0 \\ \left[\begin{array}{c|c} \sum_k \lambda_k S_k & B^T - A^T H \\ \hline B - H^T A & I_k \end{array} \right] \succeq 0 \end{array} \right\}$$

satisfies the risk bound

$$\text{Risk}[\hat{x}_{H_*} | \mathcal{X}] \leq \sqrt{\text{Opt}} \leq \sqrt{6 \ln \left(\frac{8M_*^2 K}{\text{Risk}_{\text{Opt}}^2[\mathcal{X}]} \right)} \text{Risk}_{\text{Opt}}[\mathcal{X}],$$

Numerical illustration

In these experiments

- B is $n \times n$ identity matrix,
- $n \times n$ sensing matrix A is a randomly rotated matrix with singular values λ_j , $1 \leq j \leq n$, forming a geometric progression, with $\lambda_1 = 1$ and $\lambda_n = 0.01$.
- In the first experiment the signal set \mathcal{X}_1 is an ellipsoid:

$$\mathcal{X}_1 = \{x \in \mathbb{R}^n : \sum_{j=1}^n j^2 x_j^2 \leq 1\},$$

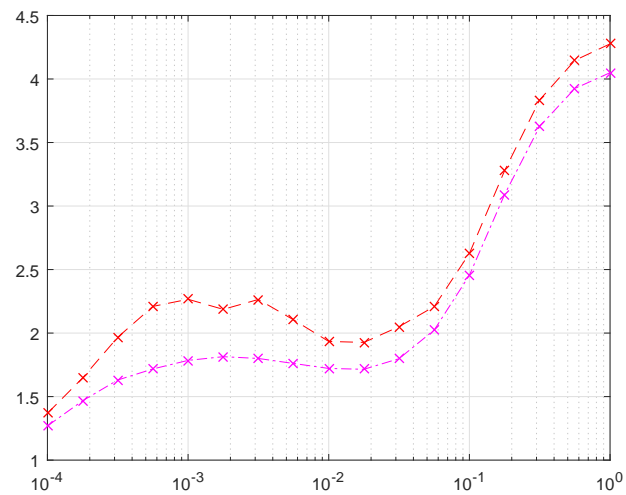
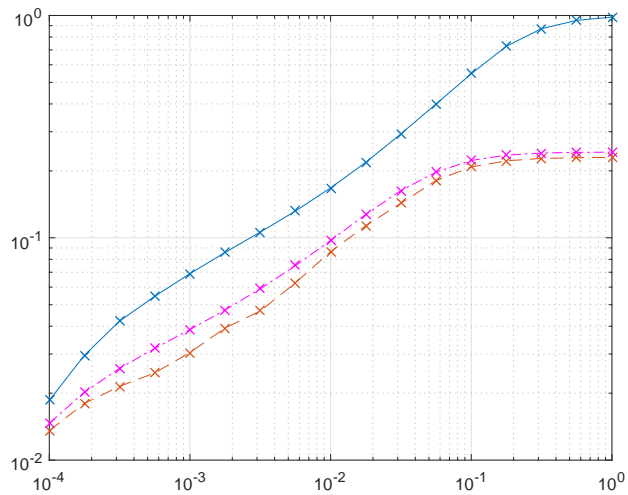
that is, $K = 1$, $S_1 = \sum_{j=1}^n j^2 e_j e_j^T$ (e_j are basic orths), and $\mathcal{T} = [0, 1]$.

Theoretical “suboptimality factor” in the interval [31.6, 73.7] in this experiment.

- In the second experiment, the signal set \mathcal{X} is the box:

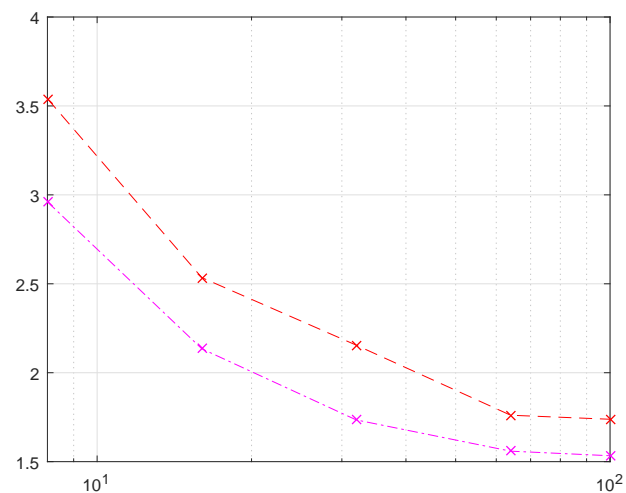
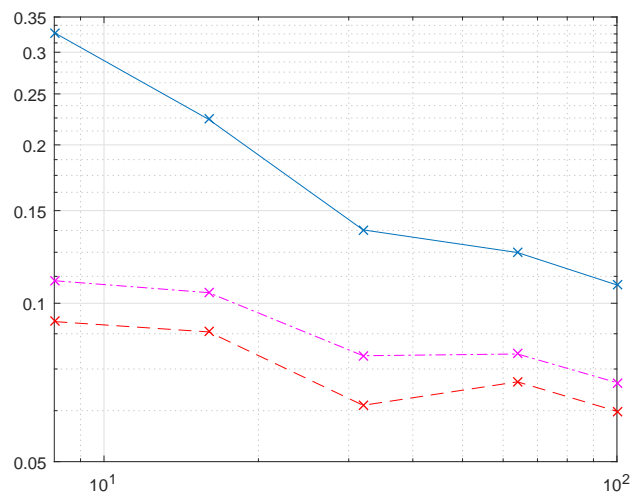
$$\mathcal{X} = \{x \in \mathbb{R}^n : j|x_j| \leq 1, 1 \leq j \leq n\} \quad [K = n, S_k = k^2 e_k e_k^T, k = 1, \dots, K, \mathcal{T} = [0, 1]^K].$$

Theoretical “suboptimality factor” in the interval [73.2, 115.4].



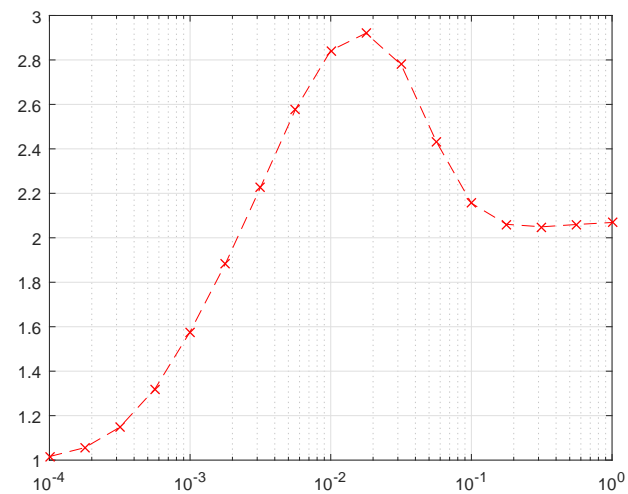
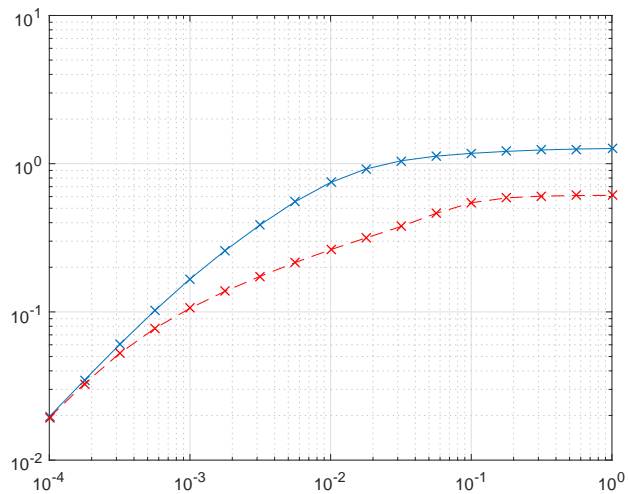
Recovery on ellipsoids: risk bounds as functions of the noise level σ , dimension $n = 32$.

Left plot: upper and lower bounds of the risk; right plot: suboptimality ratios.



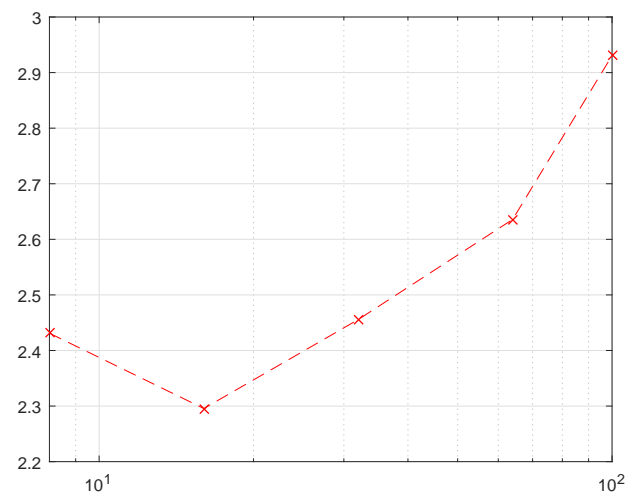
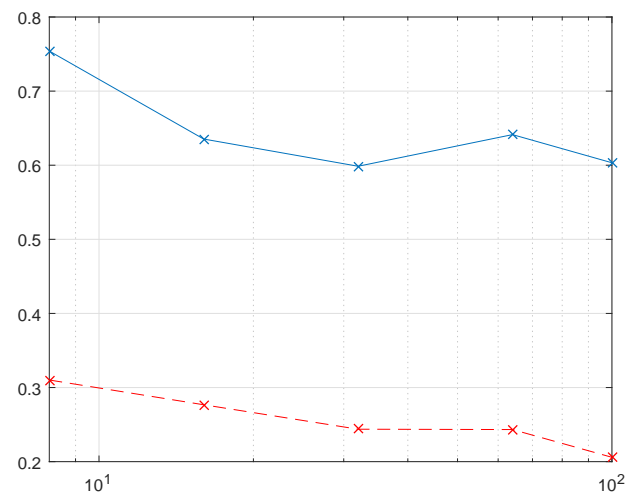
Recovery on ellipsoids: risk bounds as functions of problem dimension n , noise level $\sigma = 0.01$.

Left plot: upper and lower risk bounds; right plot: suboptimality ratios.



Recovery on a box: risk bounds as functions of the noise level σ , dimension $n = 32$.

Left plot: upper and lower bounds of the risk; right plot: suboptimality ratios.



Recovery on a box: risk bounds as functions of problem dimension n , noise level $\sigma = 0.01$.

Left plot: upper and lower risk bounds; right plot: suboptimality ratios.

Application: Inverse Heat Equation

Situation: square plate is heated at time 0 and is rest to cool; the temperature at the plate's boundary is kept 0 all the time.

Given noisy measurements of plate's temperature at time t_1 , taken along m points, we want to recover distribution of temperature at a given time t_0 , $0 < t_0 < t_1$.

The Model: the temperature field $u(t; p, q)$ evolves according to *Heat Equation*

$$\frac{\partial}{\partial t} u(t; p, q) = \left[\frac{\partial}{\partial p^2} + \frac{\partial^2}{\partial q^2} \right] u(t; p, q), \quad t \geq 0, (p, q) \in S$$

- t : time
- $S = \{(p, q) : -1 \leq p, q \leq 1\}$: the plate

with boundary conditions

$$u(t; p, q) \Big|_{(p,q) \in \partial S} \equiv 0.$$

It is convenient to represent $u(t; p, q)$ by its expansion

$$u(t; p, q) = \sum_{k,\ell} x_{k\ell}(t) \phi_k(p) \phi_\ell(q), \quad (*)$$
$$\phi_k(s) = \begin{cases} \cos(\omega_{2i-1}s), & \omega_{2i-1} = (i - 1/2)\pi, & k = 2i - 1 \\ \sin(\omega_{2i}s), & \omega_{2i} = i\pi, & k = 2i \end{cases}$$

Note:

- $\{\phi_{k\ell}(p, q) = \phi_k(p)\phi_\ell(q)\}_{k,\ell}$ form an orthonormal basis in $L_2(S)$
- $\phi_{k\ell}(\cdot)$ meet the boundary conditions $\phi_{k\ell}(p, q)|_{(p,q) \in \partial S} = 0$
- in terms of the coefficients $x_{k\ell}(t)$, the Heat Equation becomes

$$\frac{d}{dt}x_{k\ell}(t) = -(\omega_k^2 + \omega_\ell^2)x_{k\ell}(t) \Rightarrow x_{k\ell}(t) = e^{-(\omega_k^2 + \omega_\ell^2)t}x_{k\ell}(0).$$

We select integer discretization parameter N and

- restrict (*) to terms with $1 \leq k, \ell \leq 2N - 1$
- discretize the spatial variable (p, q) to reside on the grid

$$G_N = \{P_{ij} = (p_i, p_j) = (i/N - 1, j/N - 1), 1 \leq i, j \leq 2N - 1\}$$

Restricting functions $\phi_{k\ell}(\cdot, \cdot)$, $1 \leq k, \ell \leq 2N - 1$ on the grid G_N , we get orthogonal basis in $\mathbb{R}^{(2N-1) \times (2N-1)}$.

The model

- The signal x underlying observation is $x = \{x_{k\ell}(0), 1 \leq k, \ell \leq 2N - 1\} \in \mathbb{R}^{(2N-1) \times (2N-1)}$
- The observation is $\omega = A(x) + \sigma\xi \in \mathbb{R}^m$, $\xi \sim \mathcal{N}(0, I_m)$, where

$$[A(x)]_\nu = \sum_{k,\ell=1}^{2N-1} x_{k\ell} e^{-(\omega_k^2 + \omega_\ell^2)t_1} \phi_k(p_i(\nu)) \phi_\ell(p_j(\nu))$$

- $(p_i(\nu), p_j(\nu)) \in S$, $1 \leq \nu \leq m$: measurement points

- We want to recover the restriction $B(x)$ of $u(t_0; p, q)$ of some grid, say, square grid

$$G_K = \left\{ r_i = \frac{i}{K} - 1, r_j = \frac{j}{K} - 1, 1 \leq i, j \leq 2K - 1 \right\} \subset S,$$

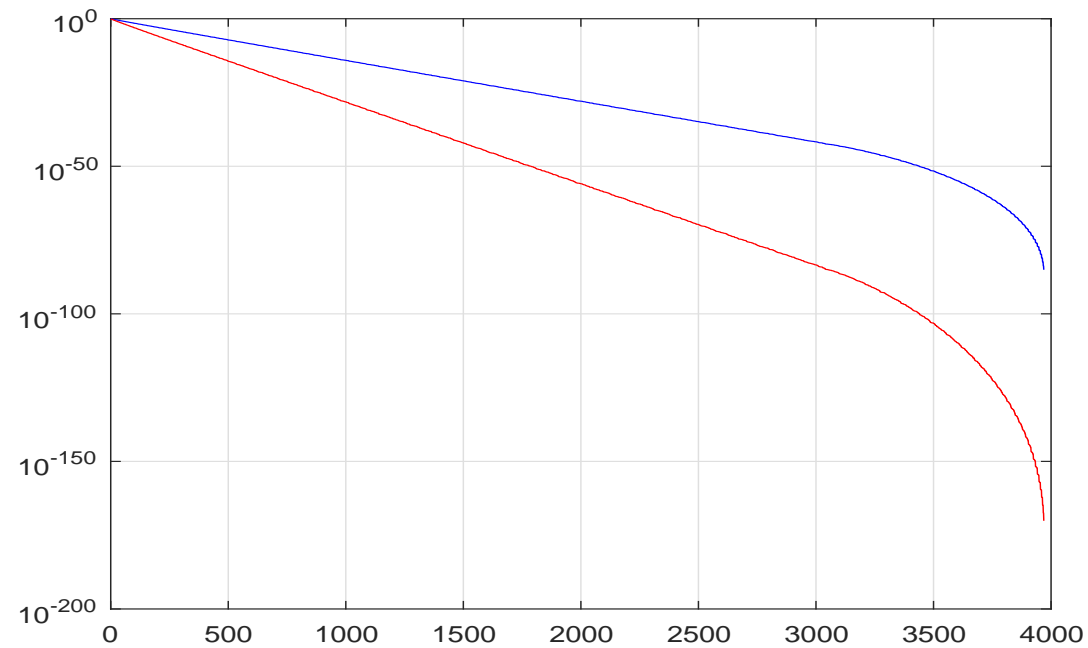
resulting in

$$[B(x)]_{ij} = \sum_{k,\ell=1}^{2N-1} x_{k\ell}(0) e^{-(\omega_k^2 + \omega_\ell^2)t_0} \phi_k(r_i) \phi_\ell(r_j)$$

- We assume that the initial distribution of temperatures $[u(0; p_i, p_j)]_{i,j=1}^{2N-1}$ satisfies $\|u\|_2 \leq R$, for some given R , implying that x resides in a centered at the origin ball

$$\mathcal{X} \subset \mathbb{R}^{(2N-1) \times (2N-1)}$$

$$u(t; p_i, p_j) = \sum_{k, \ell} e^{-(\omega_k^2 + \omega_\ell^2)t} x_{k\ell}(0) \phi_k(p_i) \phi_\ell(p_j)$$



Plots of $e^{-(\omega_k^2 + \omega_\ell^2)t}$ for $t = 0.02$ (red) and $t = 0.01$ (blue)

Bad news: contributions of high frequency (with large $\omega_k^2 + \omega_\ell^2$) components $x_{k\ell}(0)$ to $A(x)$ decrease exponentially fast with high decay rate as t_1 grows
 \Rightarrow High frequency components $x_{k\ell}(0)$ are impossible to recover from observations at time t_1 , unless t_1 is very small (red curve, $t = 0.02$)

Good news: Contributions of high frequency components $x_{k\ell}(0)$ to $B(x)$ are very small, provided t_0 is not too small
 \Rightarrow There is no need to recover well high frequency components, provided they are not huge (blue curve, $t = 0.01$)

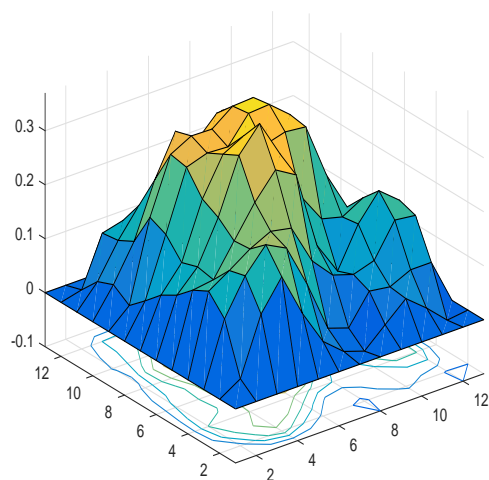
Numerical results $N = 32$, $m = 125$, $K = 6$, $t_0 = 0.01$, $t_1 = 0.02$, $\sigma = 0.001$,

$$\mathcal{X} = \{x \in \mathbb{R}^{63 \times 63} : \|\{u(0; p_i, p_j)\}_{i,j=1}^{63}\|_2 \leq 15\}.$$

- Minimax risk of optimal linear estimate: 0.1707



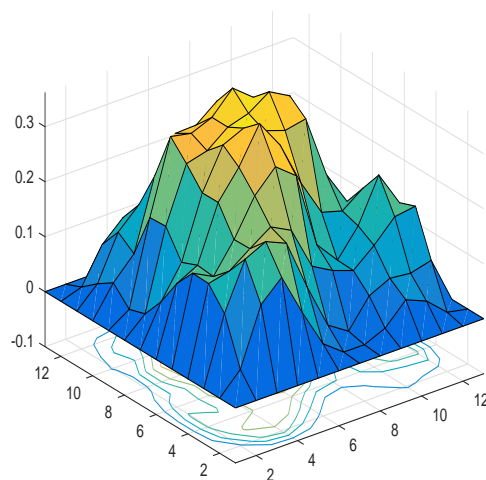
63×63 grid G_{63} and $m = 125$ measurement points



B

$$\|B\|_2 = 2.128$$

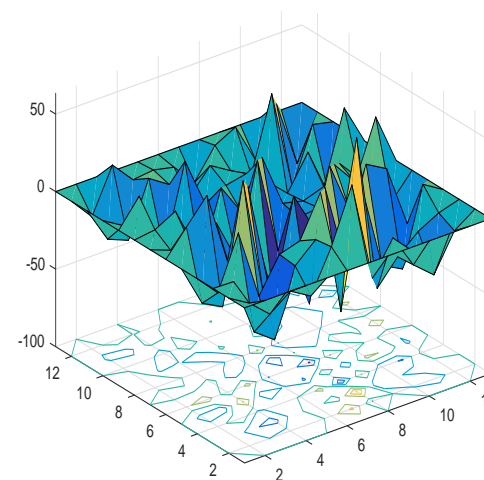
$$\|B\|_\infty = 0.433$$



\hat{B}

$$\|B - \hat{B}\|_2 = 0.1535$$

$$\|B - \hat{B}\|_\infty = 0.0313$$



\tilde{B}

$$\|B - \tilde{B}\|_2 = 266.3$$

$$\|B - \tilde{B}\|_\infty = 63.63$$

Left: true signal, center: linear recovery, right: naive LS recovery

Extension I: relative risks

When “very large” signals are allowed, one may switch from the usual risk to its *relative version* – “*S*-risk” defined as follows:

- Given a positive semidefinite “risk calibrating matrix” S we set

$$\text{Risk}_S[\hat{x}|\mathcal{X}] = \min \left\{ \sqrt{\tau} : \mathbf{E}_\xi \left\{ \|Bx - \hat{x}(Ax + \sigma\xi)\|_2^2 \right\} \leq \tau[1 + x^T Sx] \forall x \in \mathcal{X} \right\}$$

Note: setting $S = 0$ recovers the usual “plain” risk.

- Results on design of near-optimal, in terms of plain risk, linear estimates extend directly to the case of *S*-risk.

Design of near optimal linear estimate $\hat{x}_{H_*}(\omega) = H_*^T \omega$ is given by an optimal solution (H_*, τ_*, λ_*) to the convex optimization problem

$$\text{Opt} = \min_{H, \tau, \lambda} \left\{ \tau : \left[\begin{array}{c|c} \sum_k \lambda_k S_k + \tau S & B^T - A^T H \\ \hline B - H^T A & I_k \end{array} \right] \succeq 0, \sigma^2 \text{Tr}(H H^T) + \phi_{\mathcal{T}}(\lambda) \leq \tau, \lambda \geq 0 \right\}$$

For the resulting estimate, it holds

$$\text{RiskS}[\hat{x}_{H_*} | \mathcal{X}] \leq \sqrt{\text{Opt}},$$

provided ξ is zero mean with unit covariance matrix.

Near-optimality properties of the estimate \hat{x}_{H_*} remain the same as in the case of plain risk: when $\xi \sim \mathcal{N}(0, I_m)$, one has

$$\text{RiskS}[\hat{x}_{H_*} | \mathcal{X}] \leq \sqrt{6 \ln \left(\frac{8 K M_*^2}{\text{RiskS}_{\text{Opt}}^2[\mathcal{X}]} \right)} \text{RiskS}_{\text{Opt}}[\mathcal{X}],$$

where

$$M_* = \max_Q \left\{ \sqrt{\text{Tr}(B Q B^T)} : Q \succeq 0, \exists t \in \mathcal{T} : \text{Tr}(Q S_k) \leq t_k \ 1 \leq k \leq K \right\}$$

and

$$\text{RiskS}_{\text{Opt}}[\mathcal{X}] = \inf_{\hat{x}(\cdot)} \text{RiskS}[\hat{x} | \mathcal{X}].$$

In the case $\mathcal{X} = \mathbb{R}^n$, the best linear estimate is yielded by optimal solution to the convex problem

$$\text{Opt} = \min_{H, \tau} \left\{ \tau : \left[\begin{array}{c|c} \tau S & B^T - A^T H \\ \hline B - H^T A & I_k \end{array} \right] \succeq 0, \sigma^2 \text{Tr}(HH^T) \leq \tau \right\} \quad (*)$$

A feasible solution τ, H to $(*)$ gives rise to linear estimate $\hat{x}_H(\omega) = H^T \omega$ such that

$$\text{RiskS}[\hat{x}_H | \mathbb{R}^n] \leq \sqrt{\tau},$$

provided ξ is zero mean with unit covariance matrix.

Proposition. *Assume that $B \neq 0$ and $(*)$ is feasible. Then the problem is solvable, and its optimal solution Opt, H_* gives rise to linear estimate*

$$\hat{x}_{H_*}(\omega) = H_*^T \omega$$

with S -risk $\sqrt{\text{Opt}}$.

When $\xi \sim \mathcal{N}(0, I_m)$, this estimate is minimax optimal:

$$\text{RiskS}[\hat{x}_{H_*} | \mathbb{R}^n] = \sqrt{\text{Opt}} = \text{RiskS}_{\text{Opt}}[\mathbb{R}^n].$$

Application: estimating the input of a dynamical system

Let $[r; v]$ be the state of a pendulum – a 2-dimensional dynamical system satisfying

$$\dot{r} = v, \quad \dot{v} = -\nu^2 r - \kappa v + w,$$

where w is the external input. When discretizing the system, and assuming that the input is constant on the discretization interval Δ , we come to finite-difference equation

$$z_\tau = Pz_{\tau-1} + Qw_\tau, \quad \tau = 1, 2, \dots$$

where

$$P = \exp\{\Delta\vartheta\}, \quad Q = \int_0^\Delta \exp\{s\vartheta\} \begin{bmatrix} 0 \\ 1 \end{bmatrix} ds, \quad \vartheta = \begin{bmatrix} 0 & 1 \\ -\nu^2 & -\kappa \end{bmatrix}.$$

We assume that noisy observations of z_τ are available for $\tau = 1, \dots, T$, and we want to recover “signals” – collections

$$w^k = [w_{T-k+1}, \dots, w_T] = B^{(k)}x, \quad x = [z_0; w_1; \dots; w_T].$$

We want to build the best in terms of S -risk estimate.

Building linear estimate

In our design S is not fixed in advance, but satisfies $S \succeq 0$, $\text{Tr}(S) \leq 1$.

To compute the estimate $\hat{w}^k = H_*^T \omega$ of w^k we solve

$$\text{Opt}[B] = \min_{\tau, H, S} \left\{ \tau : \begin{bmatrix} \tau S & B^T - A^T H \\ B - H^T A & I_T \end{bmatrix} \succeq 0, \sigma^2 \text{Tr}(H^T H) \leq \tau, S \succeq 0, \text{Tr}(S) \leq 1 \right\}$$

with $B = B^{(k)}$.

Interpretation: recall that $\text{RiskS}[\hat{w}^k | \mathbb{R}^n] \leq \sqrt{\tau}$

$$\Leftrightarrow \mathbf{E}_{\xi \sim \mathcal{N}(0, I)} \{ \|\hat{w}^k(Ax + \sigma\xi) - B^k x\|_2^2 \} \leq \tau(1 + x^T S x) \quad \forall x.$$

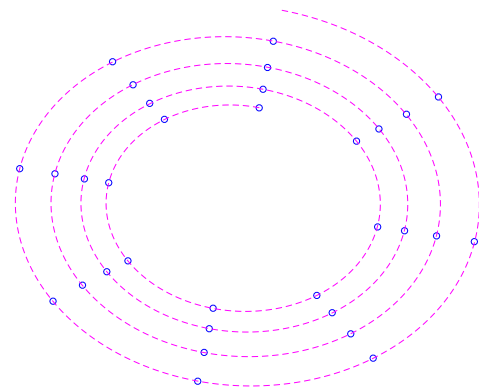
Assuming that x is random and distributed over a sphere S of radius

$$\sqrt{\dim x} = \sqrt{T + 2},$$

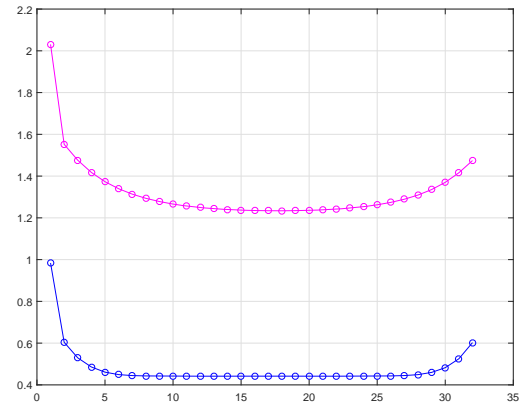
we have

$$\mathbf{E}_{x, \xi} \{ \|\hat{w}^k(Ax + \sigma\xi) - B^k x\|_2^2 \} \leq 2\tau.$$

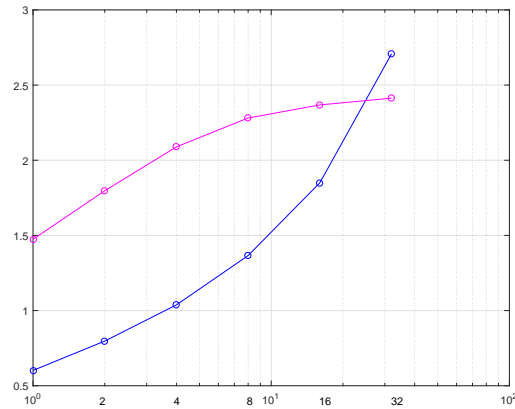
We refer to $\sqrt{\tau}$ as to *Bayesian risk* of \hat{x}^k .



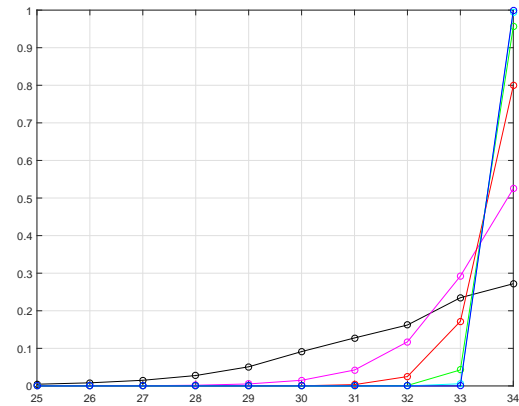
(a)



(b)



(c)



(d)

(a) free motion ($w \equiv 0$) in (r, v) -plane; (b) Bayesian and worst-case risks of recovering w_t vs. $t = 1, 2, \dots, 32$; (c) Bayesian and worst-case risks of recovering $w^k := [w_{T-k+1}; w_{T-k+2}; \dots; w_T]$ vs. k ; (d) 10 largest eigenvalues $\lambda_i(S_k)$ of S_k ($k = 32, 16, 8, 4, 2$ and 1).

Extension II: linear estimation over spectratopes

We say that a set $\mathcal{X} \subset \mathbb{R}^n$ is a *basic spectratope*, if it can be represented in the form

$$\mathcal{X} = \{x \in \mathbb{R}^n : \exists t \in \mathcal{T} : R_k^2[x] \preceq t_k I_{d_k}, 1 \leq k \leq K\}$$

where

[S₁] $R_k[x] = \sum_{i=1}^n x_i R^{ki}$ are symmetric $d_k \times d_k$ matrices linearly depending on $x \in \mathbb{R}^n$ (i.e., “matrix coefficients” R^{ki} belong to \mathbf{S}^n)

[S₂] $\mathcal{T} \in \mathbb{R}_+^K$ is a convex compact subset of \mathbb{R}_+^K which contains a positive vector and is monotone:

$$0 \leq t' \leq t \in \mathcal{T} \Rightarrow t' \in \mathcal{T}.$$

[S₃] Whenever $x \neq 0$, it holds $R_k[x] \neq 0$ for at least one $k \leq K$.

A *spectratope* is a linear image $\mathcal{Y} = PX$ of a basic spectratope.

We refer to $d = \sum_k d_k$ as *size of the spectratope* \mathcal{Y} .

Examples

[A.] Any ellitope is a spectratope.

[B.] Let L be a positive definite $d \times d$ matrix. Then the “*matrix box*”

$$\mathcal{X} = \{X \in \mathbf{S}^d : -L \preceq X \preceq L\} = \{X \in \mathbf{S}^d : R^2[X] := [L^{-1/2}XL^{-1/2}]^2 \preceq I_d\}$$

is a basic spectratope. As a result, a *bounded* set $\mathcal{X} \subset \mathbb{R}^\nu$ given by a system of “two-sided” LMI’s, specifically,

$$\mathcal{X} = \{x \in \mathbb{R}^\nu : \exists t \in \mathcal{T} : -t_k L_k \preceq S_k[x] \preceq t_l L_k, 1 \leq k \leq K\}$$

where $S_k[x]$ are symmetric $d_k \times d_k$ matrices linearly depending on x , $L_k \succ 0$ and \mathcal{T} satisfies S_2 , is a basic spectratope:

$$\mathcal{X} = \{x \in \mathbb{R}^\nu : \exists t \in \mathcal{T} : R_k^2[x] \preceq t_k I_{d_k}, k \leq K\} \quad [R_k[x] = L_k^{-1/2} S_k[x] L_k^{-1/2}]$$

Same as ellitopes, *spectratopes admit fully algorithmic calculus.*

Building linear estimate

Proposition. Consider convex optimization problem

$$\text{Opt} = \min_{H, \Lambda, \tau} \left\{ \tau : \begin{array}{l} (B - H^T A)^T (B - H^T A) \preceq \sum_k \mathcal{R}_k^*(\Lambda_k) + \tau S \\ \sigma^2 \text{Tr}(H^T H) + \phi_\tau(\lambda[\Lambda]) \leq \tau \end{array} \right\} \quad (*)$$

where $\mathcal{R}_k^*(\Lambda) : \mathbf{S}^{d_k} \rightarrow \mathbf{S}^n$ is the conjugate linear mapping,

$$[\mathcal{R}_k^*(\Lambda)]_{ij} = \frac{1}{2} \text{Tr} (\Lambda [R^{ki} R^{kj} + R^{kj} R^{ki}]), \quad 1 \leq i, j \leq n,$$

and for $\Lambda = \{\Lambda_k \in \mathbf{S}^{d_k}\}_{k \leq K}$, $\lambda[\Lambda] = [\text{Tr}[\Lambda_1]; \dots; \text{Tr}[\Lambda_K]]$.

Problem (*) is solvable, and its feasible solution (H, λ, τ) induces a linear estimate

$\hat{x}_H = H^T \omega$ of Bx , $x \in \mathcal{X}$, via observation

$$\omega = Ax + \sigma \xi, \quad \xi \sim \mathcal{N}(0, I)$$

with S -risk not exceeding $\sqrt{\tau}$.

Proposition. Let \mathcal{X} be a spectratope, and let

$$\mathcal{Q} = \{Q \in \mathbf{S}_+^n : \exists t \in \mathcal{T} : \mathcal{R}_k[Q] \preceq t_k I_{d_k}, k \leq K\}, \quad \mathcal{Q}_\rho = \rho \mathcal{Q}, \rho > 0.$$

The set \mathcal{Q} is a nonempty convex compact set containing a neighbourhood of the origin, so that the quantity

$$M_* = \sqrt{\max_{Q \in \mathcal{Q}} \text{Tr}(BQB^T)},$$

is well defined and positive.

The efficiently computable linear estimate $\hat{x}_{H_*}(\omega) = H_*^T \omega$ yielded by the optimal solution of (*) is nearly optimal in terms of S -risk:

$$\text{RiskS}[\hat{x}_{H_*} | \mathcal{X}] \leq 2 \sqrt{2 \ln \left(\frac{8DM_*^2}{\text{RiskS}_{\text{Opt}}^2[\mathcal{X}]} \right)} \text{RiskS}_{\text{Opt}}[\mathcal{X}],$$

where

$$\text{RiskS}_{\text{Opt}}[\mathcal{X}] = \inf_{\hat{x}(\cdot)} \text{RiskS}[\hat{x} | \mathcal{X}]$$

is the minimax S -risk associated with \mathcal{X} , and $d = \sum_k d_k$.