

# *Linear Recovery from Gaussian Observations*

joint work with

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[http://www2.isye.gatech.edu/~nemirovs/StatOpt\\_LN.pdf](http://www2.isye.gatech.edu/~nemirovs/StatOpt_LN.pdf)

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**Situation:** “In the nature” there exists a signal  $x$  known to belong to a given convex compact set  $\mathcal{X} \subset \mathbb{R}^n$ . We observe corrupted by noise affine image of the signal:

$$\omega = Ax + \sigma\xi \in \Omega = \mathbb{R}^m$$

- $A$ : given  $m \times n$  sensing matrix
- $\xi$ : random observation noise
- **Our goal** is to recover the image  $Bx$  of  $x$  under a given affine mapping  $B: \mathbb{R}^n \rightarrow \mathbb{R}^k$ .
- **Risk** of a candidate estimate  $\hat{x}(\cdot) : \Omega \rightarrow \mathbb{R}^k$  is defined as

$$\text{Risk}[\hat{x}|\mathcal{X}] = \sup_{x \in \mathcal{X}} \sqrt{\mathbf{E}_{\xi} \{ \|Bx - \hat{x}(Ax + \sigma\xi)\|_2^2 \}}$$

$\Rightarrow$  Risk<sup>2</sup> is the worst-case, over  $x \in \mathcal{X}$ , expected  $\|\cdot\|_2^2$  recovery error.

**Agenda:** Under appropriate assumptions on  $\mathcal{X}$ , we are to show that

- *One can build, in a computationally efficient fashion, the (nearly) best, in terms of risk, estimate from the family of linear estimates*

$$\hat{x}(\omega) = \hat{x}_H(\omega) = H^T \omega \quad [H \in \mathbb{R}^{m \times k}]$$

- *The resulting linear estimate is nearly optimal among **all** estimates, linear and nonlinear alike.*

## Linear estimation of signal in Gaussian noise

- ...
- Kuks & Olman, 1971, 1972
- Rao 1972, 1973, Pilz, 1981, 1986, ..., Drygas, 1996, Christopeit & Helmes, 1996, Arnold & Stahlecker, 2000, ...
- Pinsker 1980, Efromovich & Pinsker, 1981, 1982, Efromovich & Pinsker 1996, Golubev Levit & Tsybakov, 1996, ..., Efromovich, 2008, ...
- Donoho, Liu, McGibbon, 1990
- ...

## Risk of linear estimation

Assuming that  $\xi$  is zero mean with unit covariance matrix, we can easily compute the risk of a linear estimate  $\hat{x}_H(\omega) = H^T \omega$

$$\begin{aligned}\text{Risk}^2[\hat{x}_H|\mathcal{X}] &= \max_{x \in \mathcal{X}} \mathbf{E}_\xi \left\{ \| [B - H^T A]x - \sigma H^T \xi \|_2^2 \right\} \\ &= \max_{x \in \mathcal{X}} \left\{ \| [B - H^T A]x \|_2^2 + \sigma^2 \mathbf{E}_\xi \{ \text{Tr}(H^T \xi \xi^T H) \} \right\} \\ &= \sigma^2 \text{Tr}(H^T H) + \max_{x \in \mathcal{X}} \text{Tr}(xx^T [B^T - A^T H][B - H^T A]).\end{aligned}$$

**Note:**  $\phi$  is convex  $\Rightarrow$  building the minimum risk linear estimate reduces to solving convex minimization problem

$$\text{Opt}^P = \min_H \left[ \phi(H) = \max_{x \in \mathcal{X}} \text{Tr}(xx^T [B^T - A^T H][B - H^T A]) + \sigma^2 \text{Tr}(H^T H) \right]. \quad (*)$$

Convex function  $\phi$  is given implicitly and can be difficult to compute, making (\*) difficult as well.

**Fact:** essentially, the only cases when (\*) is known to be easy are those when

- $\mathcal{X}$  is given as a convex hull of finite set of moderate cardinality
- $\mathcal{X}$  is an ellipsoid: for  $W \in \mathbf{S}^n$  and  $S \succ 0$

$$\max_{x^T S x \leq 1} \text{Tr}(x x^T W) = \lambda_{\max}(S^{-1}W).$$

where  $\lambda_{\max}(\cdot)$  is the maximal eigenvalue.

When  $\mathcal{X}$  is a “box,” computing  $\Psi$  is NP-hard...

- When  $\Psi$  is difficult to compute, we can to replace  $\Psi$  in the design problem (\*) with an efficiently computable convex upper bound  $\Psi^+(H)$ .
- We are about to consider a family of sets  $\mathcal{X}$  – *ellitopes* – for which reasonably tight bounds  $\Psi^+$  of desired type are available.

**An ellitope** is a set  $\mathcal{X} \subset \mathbb{R}^n$  given as

$$\mathcal{X} = \{x \in \mathbb{R}^n : \exists y \in \mathbb{R}^N, t \in \mathcal{T} : x = Py, y^T S_k y \leq t_k, 1 \leq k \leq K\}$$

where

- $P$  is a given  $n \times N$  matrix (we can assume that  $P = I_n$ ),
- $S_k \succeq 0$  are positive semidefinite matrices with  $\sum_k S_k \succ 0$
- $\mathcal{T}$  is a convex compact subset of  $K$ -dimensional nonnegative orthant  $\mathbb{R}_+^K$  such that
  - $\mathcal{T}$  contains some positive vectors
  - $\mathcal{T}$  is *monotone*: if  $0 \leq t' \leq t$  and  $t \in \mathcal{T}$ , then  $t' \in \mathcal{T}$  as well.

**Note:** every *ellitope* is a symmetric w.r.t. the origin convex compact set.

## Examples

[A.] A centered at the origin ellipsoid ( $K = 1$ ,  $\mathcal{T} = [0; 1]$ )

[B.] (Bounded) intersection of  $K$  ellipsoids/elliptic cylinders centered at the origin  
( $\mathcal{T} = \{t \in \mathbb{R}^K : 0 \leq t_k \leq 1, k \leq N\}$ )

[C.] Box  $\{x \in \mathbb{R}^n : -1 \leq x_i \leq 1\}$  ( $\mathcal{T} = \{t \in \mathbb{R}^n : 0 \leq t_k \leq 1, k \leq K = n\}$ ,  $x^T S_k x = x_k^2$ )

[D.]  $\mathcal{X} = \{x \in \mathbb{R}^n : \|x\|_p \leq 1\}$  with  $p \geq 2$

( $\mathcal{T} = \{t \in \mathbb{R}_+^n : \|t\|_{p/2} \leq 1\}$ ,  $x^T S_k x = x_k^2$ ,  $k \leq K = n$ )

*Ellitopes admit fully algorithmic calculus:* if  $\mathcal{X}_i$ ,  $1 \leq i \leq I$ , are ellitopes, so are

- linear images of  $\mathcal{X}_i$
- $\mathcal{X}_1 \times \dots \times \mathcal{X}_I$
- inverse linear images of  $\mathcal{X}_i$  under linear embeddings
- $\text{Conv}(\bigcup_i \mathcal{X}_i)$
- $\mathcal{X}_1 + \dots + \mathcal{X}_I$
- $\bigcap_i \mathcal{X}_i$
- ...



## Observation

Let

$$\mathcal{X} = \{x \in \mathbb{R}^n : \exists t \in \mathcal{T} : x^T S_k x \leq t_k, 1 \leq k \leq K\}$$

be an ellitope. Given a quadratic form  $x^T W x$ ,  $W \in \mathbf{S}^n$ , one has

$$\max_{x \in \mathcal{X}} x^T W x = \max_{x \in \mathcal{X}} \text{Tr}(x x^T W) \leq \max_{Q \in \mathcal{Q}} \text{Tr}(Q W),$$

where

$$\mathcal{Q} := \{Q \in \mathbf{S}^n : Q \succeq 0, \exists t \in \mathcal{T} : \text{Tr}(Q S_k) \leq t_k, k \leq K\}.$$

We conclude that

$$\phi(H) \leq \varphi(H) := \sigma^2 \text{Tr}(H^T H) + \max_{Q \in \mathcal{Q}} \text{Tr}(Q(A^T H - B^T)(H^T A - B)),$$

and

$$\text{Risk}^2[\hat{x}_H | \mathcal{X}] \leq \min_H \varphi(H).$$

This attracts our attention to the optimization problem

$$\text{Opt}^P = \min_H \left\{ \varphi(H) = \max_{Q \in \mathcal{Q}} \underbrace{\left[ \sigma^2 \text{Tr}(H^T H) + \text{Tr}(Q(A^T H - B^T)(H^T A - B)) \right]}_{\Phi(H, Q)} \right\}. \quad (P)$$

**Note** that (P) is the primal problem

$$\min_H \left[ \max_{Q \in \mathcal{Q}} \Phi(H, Q) \right]$$

associated with the convex-concave saddle point function  $\Phi(H, Q)$ . The **dual problem** associated with  $\Phi(H, Q)$  is

$$\max_{Q \in \mathcal{Q}} \left[ \min_H \Phi(H, Q) \right],$$

that is, the problem

$$\text{Opt}^D = \max_{Q \in \mathcal{Q}} \left\{ \psi(Q) := \min_H \left[ \sigma^2 \text{Tr}(H^T H) + \text{Tr}(Q(A^T H - B^T)(H^T A - B)) \right] \right\}. \quad (D)$$

By the Sion-Kakutani theorem, (P) and (D) are solvable with equal optimal values:  $\text{Opt}^D = \text{Opt}^P = \text{Opt}$ .

Note that the minimizer of  $\Phi(\cdot, Q)$  can be easily computed:

$$H(Q) = (\sigma^2 I_m + AQA^T)^{-1} AQB^T,$$

so that

$$\psi(Q) = \text{Tr}(B[Q - QA^T(\sigma^2 I_m + AQA^T)^{-1} AQA]B^T),$$

and the dual problem reads

$$\begin{aligned} \text{Opt} = \max_{Q, t} \left\{ \text{Tr}(B[Q - QA^T(\sigma^2 I_m + AQA^T)^{-1} AQA]B^T), \right. \\ \left. Q \succeq 0, t \in \mathcal{T}, \text{Tr}(QS_k) \leq t_k, k \leq K \right\} \end{aligned} \quad (D)$$

**In fact**, both (P) and (D) can be cast as **Semidefinite Optimization problems**.

In particular, (P) can be rewritten as

$$\text{Opt} = \min_{H, \lambda} \left\{ \sigma^2 \text{Tr}(H^T H) + \phi_{\mathcal{T}}(\lambda) : \begin{bmatrix} \sum_k \lambda_k S_k & B^T - A^T H \\ B - H^T A & I_\nu \end{bmatrix} \succeq 0, \lambda \geq 0 \right\} \quad (P)$$

where  $\phi_{\mathcal{T}} : \mathbb{R}^K \rightarrow \mathbb{R}$  is the support function of  $\mathcal{T}$ :

$$\phi_{\mathcal{T}}(\lambda) = \max_{t \in \mathcal{T}} \lambda^T t.$$

Note that (P) is efficiently solvable whenever  $\mathcal{T}$  is computationally tractable.

**Bottom line:** Given matrices  $A \in \mathbb{R}^{m \times n}$ ,  $B \in \mathbb{R}^{k \times n}$  and an ellitope

$$\mathcal{X} = \{x \in \mathbb{R}^n : \exists t \in \mathcal{T} : x^T S_k x \leq t_k, 1 \leq k \leq K\} \quad (*)$$

consider the convex optimization problems

$$\text{Opt}^P = \min_H \varphi(H) \quad \text{and} \quad \text{Opt}^D = \max_{Q \in \mathcal{Q}} \psi(Q),$$

where  $\mathcal{Q} := \{Q \in \mathbf{S}^n : Q \succeq 0, \exists t \in \mathcal{T} : \text{Tr}(Q S_k) \leq t_k, k \leq K\}$ .

- The optimal values of two problems coincide,  $\text{Opt}^P = \text{Opt}^D = \text{Opt}$ .
- When noise  $\xi$  satisfies  $\mathbf{E}\{\xi\} = 0$ , and  $\mathbf{E}\{\xi\xi^T\} = I_m$ , the risk of the linear estimate  $\hat{x}_{H_*}(\cdot)$  induced by the optimal solution  $H_*$  to the problem (this solution clearly exists provided  $\sigma > 0$ ) satisfies the risk bound

$$\text{Risk}[\hat{x}_{H_*} | \mathcal{X}] \leq \sqrt{\text{Opt}}.$$

## Bayesian risks

- *Minimax risk*  $\text{Risk}[\hat{x}|\mathcal{X}]$  is defined as the *worst*, over the signals of interest, performance of  $\hat{x}(\cdot)$
- *Bayesian risk* is the *average performance*, with the average taken over some *prior* probability distribution on the signals.

For the problem of  $\|\cdot\|_2$ -recovering  $Bx$  via noisy observation

$$\omega = Ax + \sigma\xi, \quad \xi \sim P$$

this alternative reads as follows:

- (!) Given a probability distribution  $\pi$  of signal  $x \in \mathbb{R}^n$ , find an estimate  $\hat{x}(\cdot)$  which minimizes

$$\text{Risk}^2(\hat{x}|\pi) := \int_{\pi} \left\{ \int_{\mathbb{R}^m} \|Bx - \hat{x}(Ax + \sigma\xi)\|_2^2 P(d\xi) \right\} \pi(dx)$$

- the average, over the distribution  $\pi$  of signals  $x$ , of expected  $\|\cdot\|_2^2$  recovery error of  $Bx$  via observation  $Ax + \sigma\xi$ .

Let  $P_{x,\omega}$  be the induced by  $\pi$  and  $P_\xi$  *joint distribution* of  $(x, \omega = Ax + \sigma\xi)$  on  $\mathbb{R}_x^n \times \mathbb{R}_\omega^m$ .  $P_{x,\omega}$  gives rise to

- *marginal distribution  $P_\omega$  of  $\omega$ ,*
- *conditional distribution  $P_{x|\omega}$  of  $x$  given  $\omega$ .*

We have

$$\begin{aligned} \text{Risk}^2(\hat{x}|\pi) &= \int_{\mathbb{R}_x^n \times \mathbb{R}_\omega^m} \|Bx - \hat{x}(\omega)\|_2^2 P_{x,\omega}(dx, d\omega) \\ &= \int_{\mathbb{R}_\omega^m} \left\{ \int_{\mathbb{R}_x^n} \|Bx - \hat{x}(\omega)\|_2^2 P_{x|\omega}(dx) \right\} P_\omega(d\omega) \end{aligned}$$

Assuming that the probability distribution  $\pi$  possesses finite second moments, one has

$$\min_{\hat{x}(\cdot)} \int_{\mathbb{R}_x^n \times \mathbb{R}_\omega^m} \|Bx - \hat{x}(\omega)\|_2^2 P_{x,\omega}(dx) = \int_{\mathbb{R}_x^n \times \mathbb{R}_\omega^m} \|Bx - \hat{x}_*(\omega)\|_2^2 P_{x,\omega}(dx),$$

where

$$\hat{x}_*(\omega) = \int_{\mathbb{R}_x^n} Bx P_{x|\omega}(dx).$$

**Corollary [Gauss-Markov theorem]:** Let  $x \in \mathbb{R}^n$  and  $\xi \in \mathbb{R}^m$  be independent zero-mean Gaussian random vectors. Assuming  $\sigma > 0$  and the covariance matrix of  $\xi$  to be positive definite,

- conditional, given  $\omega$ , distribution of  $x$  is normal, so that the conditional expectation  $\hat{x}_*(\omega)$  is a linear function of  $\omega$ ,
- as a result, an optimal solution  $\hat{x}_*(\cdot)$  to the risk minimization problem

$$\min_{\hat{x}(\cdot)} \mathbf{E}_{x,\xi} \left\{ \|Bx - \hat{x}(Ax + \sigma\xi)\|_2^2 \right\}$$

exists and is a linear function of  $\omega = Ax + \sigma\xi$ .

In particular, when  $\xi \sim \mathcal{N}(0, I_m)$  and  $x \sim \mathcal{N}(0, Q)$ , one has

$$\begin{aligned} \hat{x}_*(\omega) &= [\sigma^2 I_m + AQA^T]^{-1} AQB^T \omega \\ \text{Risk}^2(\hat{x}_* | \mathcal{N}(0, Q)) &= \text{Tr}(B[Q - QA^T[\sigma^2 I_m + AQA^T]^{-1}AQ]B^T) \end{aligned}$$

## Course of actions (Pinsker's program)

- Let  $\mathcal{N}(0, Q)$  be a Gaussian prior for the signal  $x$  which “sits on  $\mathcal{X}$  with high probability.” Then by the Gauss-Markov theorem the (“slightly reduced”) quantity

$$\varphi(Q) = \text{Tr}(B[Q - QA^T[\sigma^2 I_m + AQA^T]^{-1}AQ]B^T)$$

would be a lower bound on  $\text{Risk}_{\text{Opt}}^2$ .

- Note that  $\mathbf{E}_{\eta \sim \mathcal{N}(0, Q)}\{\eta^T S \eta\} = \text{Tr}(SQ)$ . Thus, selecting  $Q \succeq 0$  according to

$$\exists t \in \mathcal{T} : \text{Tr}(QS_k) \leq t_k, k \leq K$$

we ensure that  $\eta \sim \mathcal{N}(0, Q)$  sits in  $\mathcal{X}$  “on average.”

Imposing on  $Q \succeq 0$  restriction

$$\exists t \in \mathcal{T} : \text{Tr}(QS_k) \leq \rho t_k, k \leq K, \quad [\rho > 0]$$

we enforce  $\eta \sim \mathcal{N}(0, Q)$  to take values in  $\mathcal{X}$  with probability controlled by  $\rho$  and approaching 1 as  $\rho \rightarrow +0$ .



- The above considerations give rise to parametric optimization problem

$$\text{Opt}_*(\rho) = \max_{Q \succeq 0} \{\varphi(Q) : \exists t \in \mathcal{T} : \text{Tr}(QS_k) \leq \rho t_k, 1 \leq k \leq K\} \quad (P_\rho)$$

We *may expect* that for small  $\rho$  a “slightly corrected”  $\text{Opt}_*(\rho)$  is a lower bound on  $\text{Risk}_{\text{Opt}}^2$ .

- As we have just seen,  $\text{Opt}_*(1) = \text{Opt}$  (!). Since the optimal value of the (concave) optimization problem  $(P_\rho)$  is a concave function of  $\rho$ , we have

$$\text{Opt}_*(\rho) \geq \rho \text{Opt}, \quad 0 < \rho < 1.$$

Now, all we need is a simple result as follows:

**Lemma** Let  $S$  and  $Q$  be positive semidefinite  $n \times n$  matrices with  $\rho := \text{Tr}(SQ) \leq 1$ , and let  $\eta \sim \mathcal{N}(0, Q)$ . Then

$$\text{Prob} \{ \eta^T S \eta > 1 \} \leq e^{-\frac{1-\rho+\rho \ln(\rho)}{2\rho}}$$

We arrive at the following

**Theorem.** Let us associate with ellitope  $\mathcal{X} = \{x \in \mathbb{R}^n : \exists t \in \mathcal{T} : x^T S_k x \leq t_k, k \leq K\}$  the convex compact set

$$\mathcal{Q} = \{Q \in \mathbf{S}^n : Q \succeq 0, \exists t \in \mathcal{T} : \text{Tr}(QS_k) \leq t_k, k \leq K\},$$

and the quantity

$$M_* = \max_{Q \in \mathcal{Q}} \sqrt{\text{Tr}(BQB^T)}.$$

Then the linear estimate  $\hat{x}_{H_*}(\omega) = H_*^T \omega$  of  $Bx$ ,  $x \in \mathcal{X}$ , via observation  $\omega = Ax + \sigma\xi$ ,  $\xi \sim \mathcal{N}(0, I_m)$ , given by the optimal solution  $H_*$  to the convex optimization problem

$$\text{Opt} = \min_{H, \lambda} \left\{ \phi_{\mathcal{T}}(\lambda) + \sigma^2 \text{Tr}(HH^T) : \begin{array}{l} \lambda \geq 0 \\ \left[ \begin{array}{c|c} \sum_k \lambda_k S_k & B^T - A^T H \\ \hline B - H^T A & I_k \end{array} \right] \succeq 0 \end{array} \right\}$$

satisfies the risk bound

$$\text{Risk}[\hat{x}_{H_*} | \mathcal{X}] \leq \sqrt{\text{Opt}} \leq \sqrt{6 \ln \left( \frac{8M_*^2 K}{\text{Risk}_{\text{Opt}}^2[\mathcal{X}]} \right)} \text{Risk}_{\text{Opt}}[\mathcal{X}],$$

## Numerical illustration

In these experiments

- $B$  is  $n \times n$  identity matrix,
- $n \times n$  sensing matrix  $A$  is a randomly rotated matrix with singular values  $\lambda_j$ ,  $1 \leq j \leq n$ , forming a geometric progression, with  $\lambda_1 = 1$  and  $\lambda_n = 0.01$ .
- In the first experiment the signal set  $\mathcal{X}_1$  is an ellipsoid:

$$\mathcal{X}_1 = \{x \in \mathbb{R}^n : \sum_{j=1}^n j^2 x_j^2 \leq 1\},$$

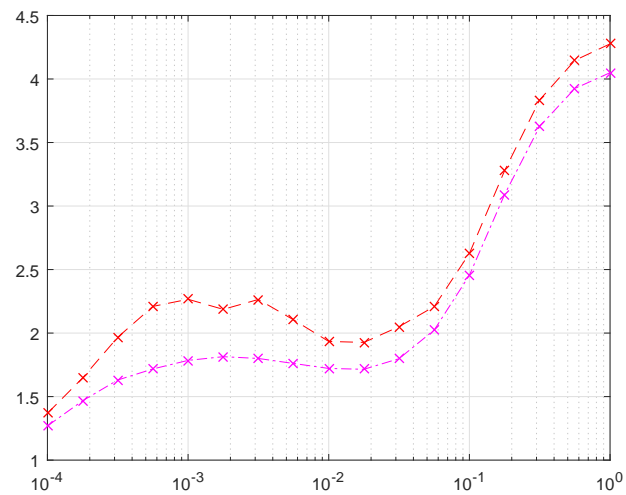
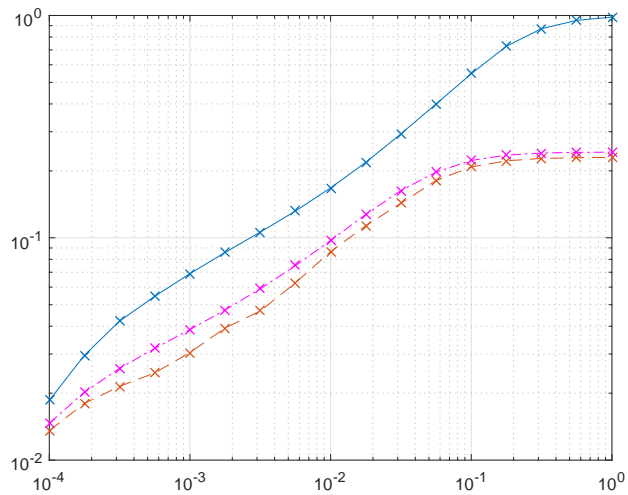
that is,  $K = 1$ ,  $S_1 = \sum_{j=1}^n j^2 e_j e_j^T$  ( $e_j$  are basic orths), and  $\mathcal{T} = [0, 1]$ .

*Theoretical “suboptimality factor” in the interval [31.6, 73.7] in this experiment.*

- In the second experiment, the signal set  $\mathcal{X}$  is the box:

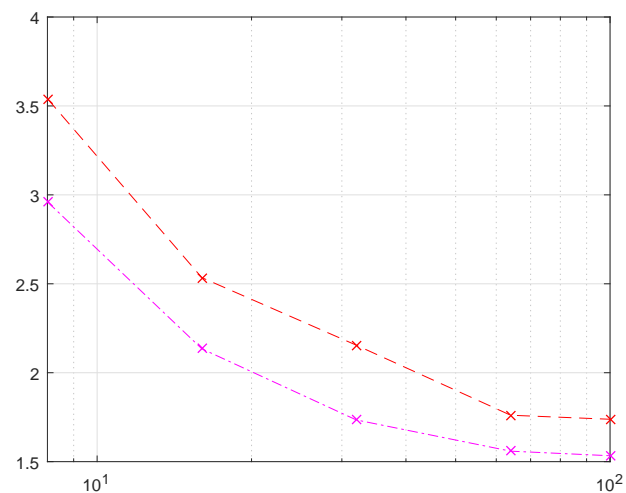
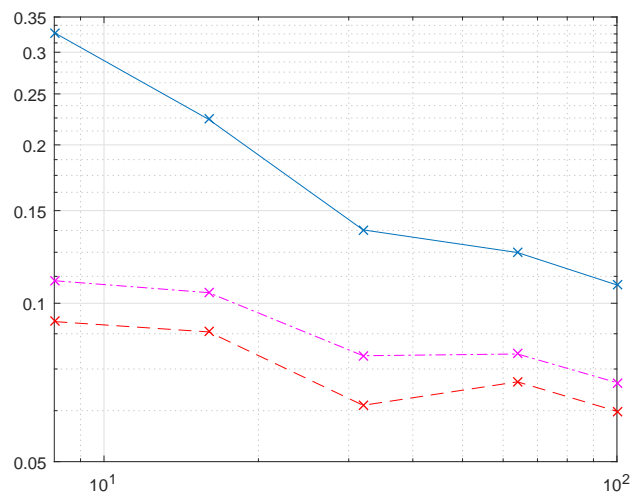
$$\mathcal{X} = \{x \in \mathbb{R}^n : |x_j| \leq 1, 1 \leq j \leq n\} \quad [K = n, S_k = k^2 e_k e_k^T, k = 1, \dots, K, \mathcal{T} = [0, 1]^K].$$

*Theoretical “suboptimality factor” in the interval [73.2, 115.4].*



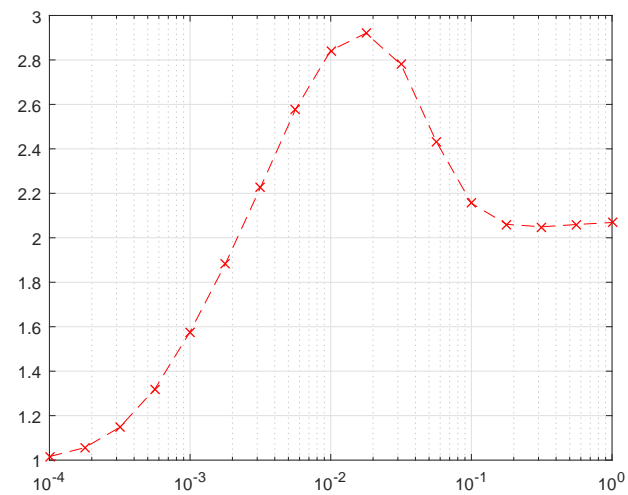
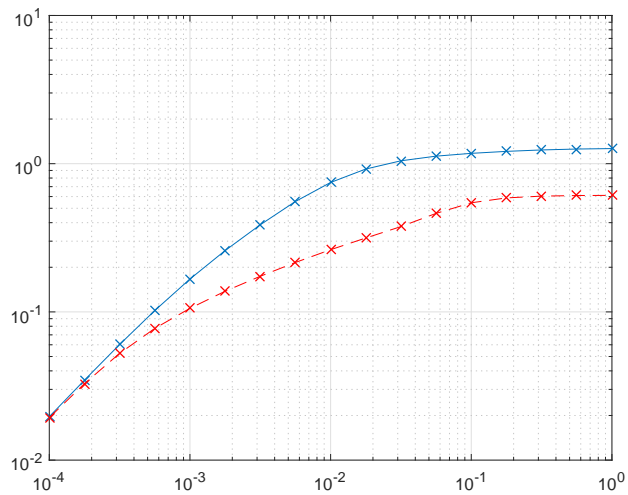
*Recovery on ellipsoids:* risk bounds as functions of the noise level  $\sigma$ , dimension  $n = 32$ .

Left plot: upper and lower bounds of the risk; right plot: suboptimality ratios.



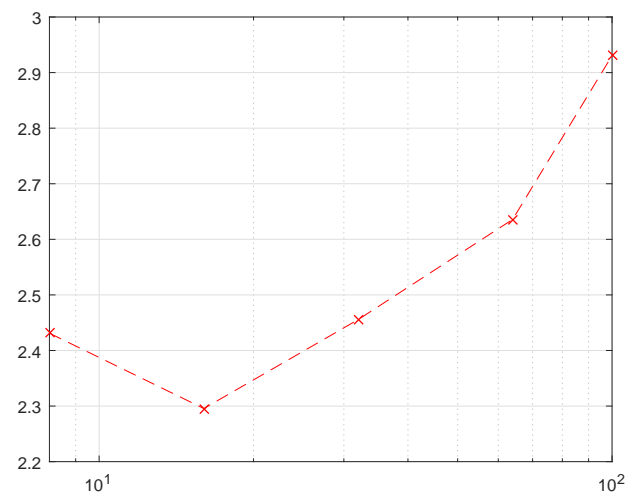
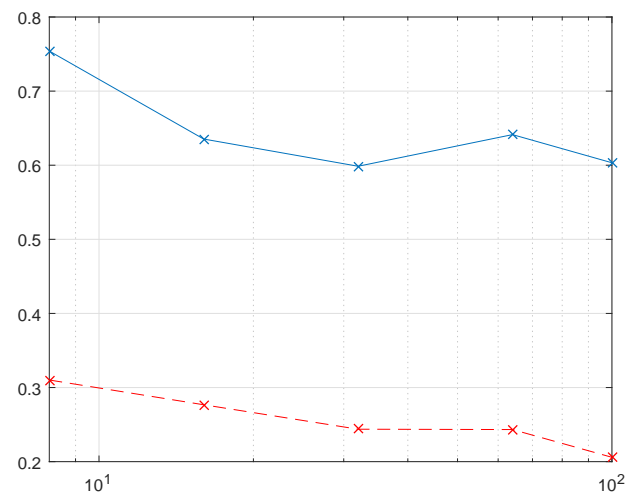
*Recovery on ellipsoids:* risk bounds as functions of problem dimension  $n$ , noise level  $\sigma = 0.01$ .

Left plot: upper and lower risk bounds; right plot: suboptimality ratios.



*Recovery on a box:* risk bounds as functions of the noise level  $\sigma$ , dimension  $n = 32$ .

Left plot: upper and lower bounds of the risk; right plot: suboptimality ratios.



*Recovery on a box:* risk bounds as functions of problem dimension  $n$ , noise level  $\sigma = 0.01$ .

Left plot: upper and lower risk bounds; right plot: suboptimality ratios.

## Application: Inverse Heat Equation

**Situation:** square plate is heated at time 0 and is rest to cool; the temperature at the plate's boundary is kept 0 all the time.

Given noisy measurements of plate's temperature at time  $t_1$ , taken along  $m$  points, we want to recover distribution of temperature at a given time  $t_0$ ,  $0 < t_0 < t_1$ .

**The Model:** the temperature field  $u(t; p, q)$  evolves according to *Heat Equation*

$$\frac{\partial}{\partial t} u(t; p, q) = \left[ \frac{\partial}{\partial p^2} + \frac{\partial^2}{\partial q^2} \right] u(t; p, q), \quad t \geq 0, (p, q) \in S$$

- $t$ : time
- $S = \{(p, q) : -1 \leq p, q \leq 1\}$ : the plate

with boundary conditions

$$u(t; p, q) \Big|_{(p,q) \in \partial S} \equiv 0.$$

It is convenient to represent  $u(t; p, q)$  by its expansion

$$u(t; p, q) = \sum_{k,\ell} x_{k\ell}(t) \phi_k(p) \phi_\ell(q), \quad (*)$$
$$\phi_k(s) = \begin{cases} \cos(\omega_{2i-1}s), & \omega_{2i-1} = (i - 1/2)\pi, & k = 2i - 1 \\ \sin(\omega_{2i}s), & \omega_{2i} = i\pi, & k = 2i \end{cases}$$

## Note:

- $\{\phi_{k\ell}(p, q) = \phi_k(p)\phi_\ell(q)\}_{k,\ell}$  form an orthonormal basis in  $L_2(S)$
- $\phi_{k\ell}(\cdot)$  meet the boundary conditions  $\phi_{k\ell}(p, q)|_{(p,q) \in \partial S} = 0$
- in terms of the coefficients  $x_{k\ell}(t)$ , the Heat Equation becomes

$$\frac{d}{dt}x_{k\ell}(t) = -(\omega_k^2 + \omega_\ell^2)x_{k\ell}(t) \Rightarrow x_{k\ell}(t) = e^{-(\omega_k^2 + \omega_\ell^2)t}x_{k\ell}(0).$$

We select integer discretization parameter  $N$  and

- restrict (\*) to terms with  $1 \leq k, \ell \leq 2N - 1$
- discretize the spatial variable  $(p, q)$  to reside on the grid

$$G_N = \{P_{ij} = (p_i, p_j) = (i/N - 1, j/N - 1), 1 \leq i, j \leq 2N - 1\}$$

Restricting functions  $\phi_{k\ell}(\cdot, \cdot)$ ,  $1 \leq k, \ell \leq 2N - 1$  on the grid  $G_N$ , we get orthogonal basis in  $\mathbb{R}^{(2N-1) \times (2N-1)}$ .

## The model

- The signal  $x$  underlying observation is  $x = \{x_{k\ell}(0), 1 \leq k, \ell \leq 2N - 1\} \in \mathbb{R}^{(2N-1) \times (2N-1)}$
- The observation is  $\omega = A(x) + \sigma\xi \in \mathbb{R}^m$ ,  $\xi \sim \mathcal{N}(0, I_m)$ , where

$$[A(x)]_\nu = \sum_{k,\ell=1}^{2N-1} x_{k\ell} e^{-(\omega_k^2 + \omega_\ell^2)t_1} \phi_k(p_i(\nu)) \phi_\ell(p_j(\nu))$$

- $(p_i(\nu), p_j(\nu)) \in S$ ,  $1 \leq \nu \leq m$ : measurement points

- We want to recover the restriction  $B(x)$  of  $u(t_0; p, q)$  of some grid, say, square grid

$$G_K = \left\{ r_i = \frac{i}{K} - 1, r_j = \frac{j}{K} - 1, 1 \leq i, j \leq 2K - 1 \right\} \subset S,$$

resulting in

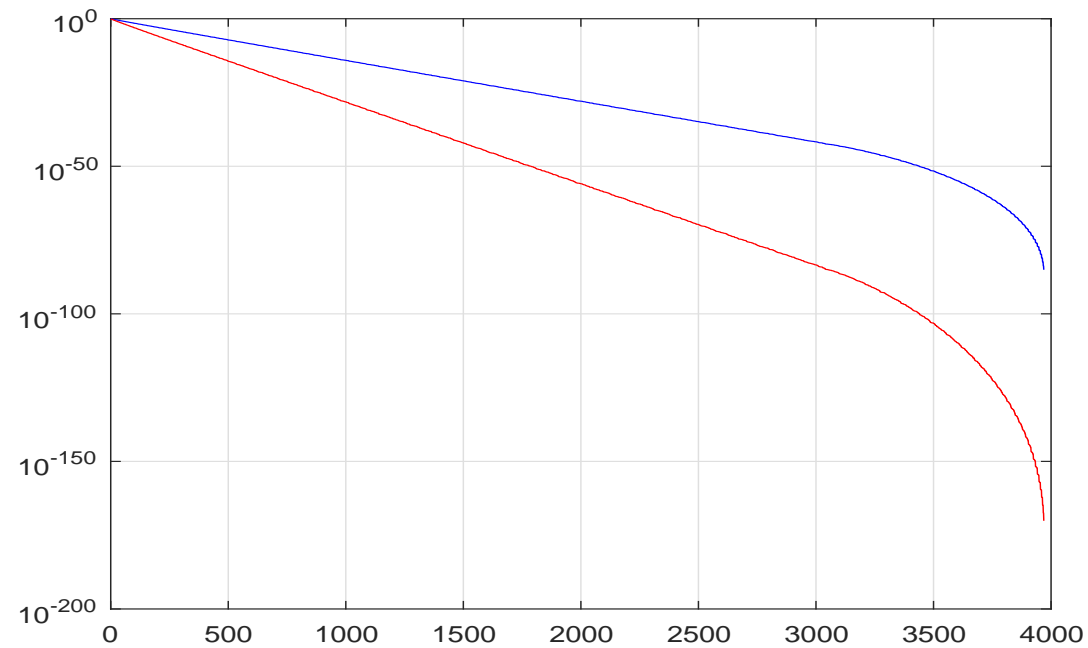
$$[B(x)]_{ij} = \sum_{k,\ell=1}^{2N-1} x_{k\ell}(0) e^{-(\omega_k^2 + \omega_\ell^2)t_0} \phi_k(r_i) \phi_\ell(r_j)$$

- We assume that the initial distribution of temperatures  $[u(0; p_i, p_j)]_{i,j=1}^{2N-1}$  satisfies  $\|u\|_2 \leq R$ , for some given  $R$ , implying that  $x$  resides in a centered at the origin ball

$$\mathcal{X} \subset \mathbb{R}^{(2N-1) \times (2N-1)}$$



$$u(t; p_i, p_j) = \sum_{k, \ell} e^{-(\omega_k^2 + \omega_\ell^2)t} x_{k\ell}(0) \phi_k(p_i) \phi_\ell(p_j)$$



Plots of  $e^{-(\omega_k^2 + \omega_\ell^2)t}$  for  $t = 0.02$  (red) and  $t = 0.01$  (blue)

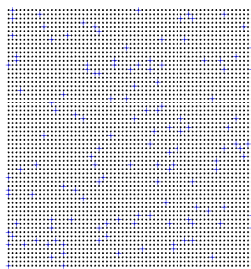
**Bad news:** contributions of high frequency (with large  $\omega_k^2 + \omega_\ell^2$ ) components  $x_{k\ell}(0)$  to  $A(x)$  decrease exponentially fast with high decay rate as  $t_1$  grows  
 $\Rightarrow$  High frequency components  $x_{k\ell}(0)$  are impossible to recover from observations at time  $t_1$ , unless  $t_1$  is very small (red curve,  $t = 0.02$ )

**Good news:** Contributions of high frequency components  $x_{k\ell}(0)$  to  $B(x)$  are very small, provided  $t_0$  is not too small  
 $\Rightarrow$  There is no need to recover well high frequency components, provided they are not huge (blue curve,  $t = 0.01$ )

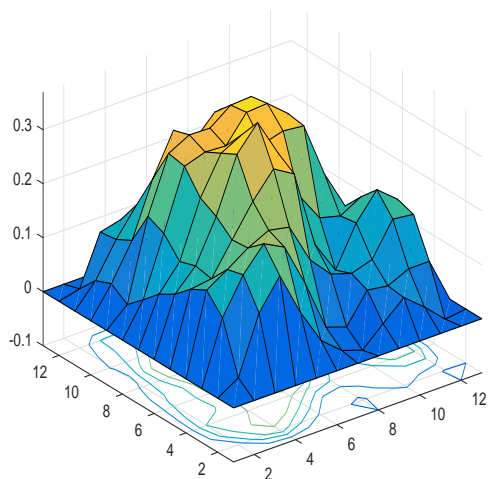
**Numerical results**  $N = 32$ ,  $m = 125$ ,  $K = 6$ ,  $t_0 = 0.01$ ,  $t_1 = 0.02$ ,  $\sigma = 0.001$ ,

$$\mathcal{X} = \{x \in \mathbb{R}^{63 \times 63} : \|\{u(0; p_i, p_j)\}_{i,j=1}^{63}\|_2 \leq 15\}.$$

- Minimax risk of optimal linear estimate: 0.1707



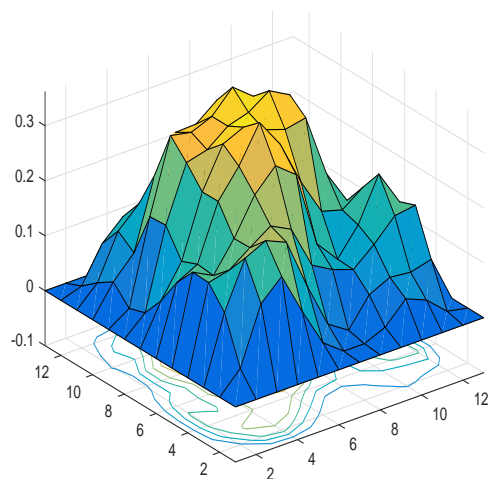
63 × 63 grid  $G_{63}$  and  $m = 125$  measurement points



$B$

$$\|B\|_2 = 2.128$$

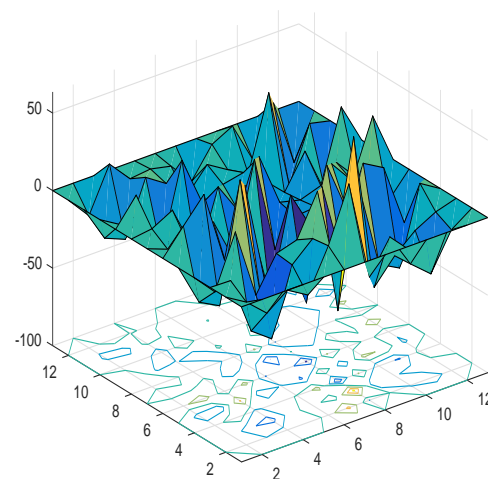
$$\|B\|_\infty = 0.433$$



$\hat{B}$

$$\|B - \hat{B}\|_2 = 0.1535$$

$$\|B - \hat{B}\|_\infty = 0.0313$$



$\tilde{B}$

$$\|B - \tilde{B}\|_2 = 266.3$$

$$\|B - \tilde{B}\|_\infty = 63.63$$

Left: true signal, center: linear recovery, right: naive LS recovery

## Extension I: relative risks

When “very large” signals are allowed, one may switch from the usual risk to its *relative version* – “*S*-risk” defined as follows:

- Given a positive semidefinite “risk calibrating matrix”  $S$  we set

$$\text{Risk}_S[\hat{x}|\mathcal{X}] = \min \left\{ \sqrt{\tau} : \mathbf{E}_\xi \left\{ \|Bx - \hat{x}(Ax + \sigma\xi)\|_2^2 \right\} \leq \tau[1 + x^T Sx] \forall x \in \mathcal{X} \right\}$$

**Note:** setting  $S = 0$  recovers the usual “plain” risk.

- Results on design of near-optimal, in terms of plain risk, linear estimates extend directly to the case of *S*-risk.

**Design** of near optimal linear estimate  $\hat{x}_{H_*}(\omega) = H_*^T \omega$  is given by an optimal solution  $(H_*, \tau_*, \lambda_*)$  to the convex optimization problem

$$\text{Opt} = \min_{H, \tau, \lambda} \left\{ \tau : \left[ \begin{array}{c|c} \sum_k \lambda_k S_k + \tau S & B^T - A^T H \\ \hline B - H^T A & I_k \end{array} \right] \succeq 0, \sigma^2 \text{Tr}(H H^T) + \phi_{\mathcal{T}}(\lambda) \leq \tau, \lambda \geq 0 \right\}$$

For the resulting estimate, it holds

$$\text{RiskS}[\hat{x}_{H_*} | \mathcal{X}] \leq \sqrt{\text{Opt}},$$

provided  $\xi$  is zero mean with unit covariance matrix.

**Near-optimality** properties of the estimate  $\hat{x}_{H_*}$  remain the same as in the case of plain risk: when  $\xi \sim \mathcal{N}(0, I_m)$ , one has

$$\text{RiskS}[\hat{x}_{H_*} | \mathcal{X}] \leq \sqrt{6 \ln \left( \frac{8KM_*^2}{\text{RiskS}_{\text{Opt}}^2[\mathcal{X}]} \right)} \text{RiskS}_{\text{Opt}}[\mathcal{X}],$$

where

$$M_* = \max_Q \left\{ \sqrt{\text{Tr}(BQB^T)} : Q \succeq 0, \exists t \in \mathcal{T} : \text{Tr}(QS_k) \leq t_k \ 1 \leq k \leq K \right\}$$

and

$$\text{RiskS}_{\text{Opt}}[\mathcal{X}] = \inf_{\hat{x}(\cdot)} \text{RiskS}[\hat{x} | \mathcal{X}].$$

**In the case**  $\mathcal{X} = \mathbb{R}^n$ , the best linear estimate is yielded by optimal solution to the convex problem

$$\text{Opt} = \min_{H, \tau} \left\{ \tau : \left[ \begin{array}{c|c} \tau S & B^T - A^T H \\ \hline B - H^T A & I_k \end{array} \right] \succeq 0, \sigma^2 \text{Tr}(HH^T) \leq \tau \right\} \quad (*)$$

A feasible solution  $\tau, H$  to  $(*)$  gives rise to linear estimate  $\hat{x}_H(\omega) = H^T \omega$  such that

$$\text{RiskS}[\hat{x}_H | \mathbb{R}^n] \leq \sqrt{\tau},$$

provided  $\xi$  is zero mean with unit covariance matrix.

**Proposition.** *Assume that  $B \neq 0$  and  $(*)$  is feasible. Then the problem is solvable, and its optimal solution  $\text{Opt}, H_*$  gives rise to linear estimate*

$$\hat{x}_{H_*}(\omega) = H_*^T \omega$$

with  $S$ -risk  $\sqrt{\text{Opt}}$ .

When  $\xi \sim \mathcal{N}(0, I_m)$ , this estimate is minimax optimal:

$$\text{RiskS}[\hat{x}_{H_*} | \mathbb{R}^n] = \sqrt{\text{Opt}} = \text{RiskS}_{\text{Opt}}[\mathbb{R}^n].$$

## Application: estimating the input of a dynamical system

Let  $[r; v]$  be the state of a pendulum – a 2-dimensional dynamical system satisfying

$$\dot{r} = v, \quad \dot{v} = -\nu^2 r - \kappa v + w,$$

where  $w$  is the external input. When discretizing the system, and assuming that the input is constant on the discretization interval  $\Delta$ , we come to finite-difference equation

$$z_\tau = Pz_{\tau-1} + Qw_\tau, \quad \tau = 1, 2, \dots$$

where

$$P = \exp\{\Delta\vartheta\}, \quad Q = \int_0^\Delta \exp\{s\vartheta\} \begin{bmatrix} 0 \\ 1 \end{bmatrix} ds, \quad \vartheta = \begin{bmatrix} 0 & 1 \\ -\nu^2 & -\kappa \end{bmatrix}.$$

We assume that noisy observations of  $z_\tau$  are available for  $\tau = 1, \dots, T$ , and we want to recover “signals” – collections

$$w^k = [w_{T-k+1}, \dots, w_T] = B^{(k)}x, \quad x = [z_0; w_1; \dots; w_T].$$

We want to build the best in terms of  $S$ -risk estimate.

## Building linear estimate

In our design  $S$  is not fixed in advance, but satisfies  $S \succeq 0$ ,  $\text{Tr}(S) \leq 1$ .

To compute the estimate  $\hat{w}^k = H_*^T \omega$  of  $w^k$  we solve

$$\text{Opt}[B] = \min_{\tau, H, S} \left\{ \tau : \begin{bmatrix} \tau S & B^T - A^T H \\ B - H^T A & I_T \end{bmatrix} \succeq 0, \sigma^2 \text{Tr}(H^T H) \leq \tau, S \succeq 0, \text{Tr}(S) \leq 1 \right\}$$

with  $B = B^{(k)}$ .

**Interpretation:** recall that  $\text{RiskS}[\hat{w}^k | \mathbb{R}^n] \leq \sqrt{\tau}$

$$\Leftrightarrow \mathbf{E}_{\xi \sim \mathcal{N}(0, I)} \{ \|\hat{w}^k(Ax + \sigma\xi) - B^k x\|_2^2 \} \leq \tau(1 + x^T S x) \quad \forall x.$$

Assuming that  $x$  is random and distributed over a sphere  $S$  of radius

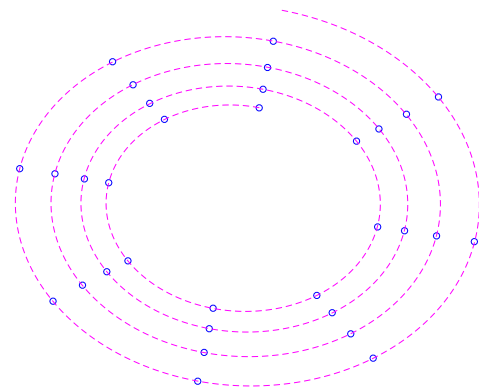
$$\sqrt{\dim x} = \sqrt{T + 2},$$

we have

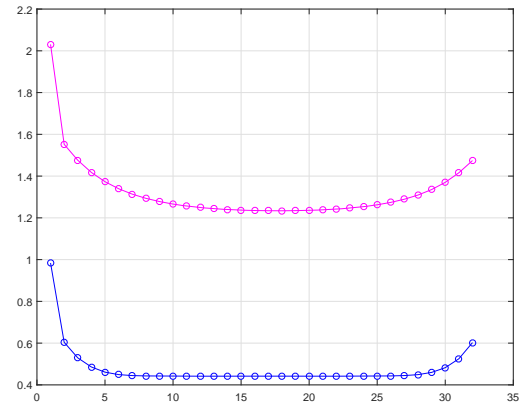
$$\mathbf{E}_{x, \xi} \{ \|\hat{w}^k(Ax + \sigma\xi) - B^k x\|_2^2 \} \leq 2\tau.$$

We refer to  $\sqrt{\tau}$  as to *Bayesian risk* of  $\hat{x}^k$ .

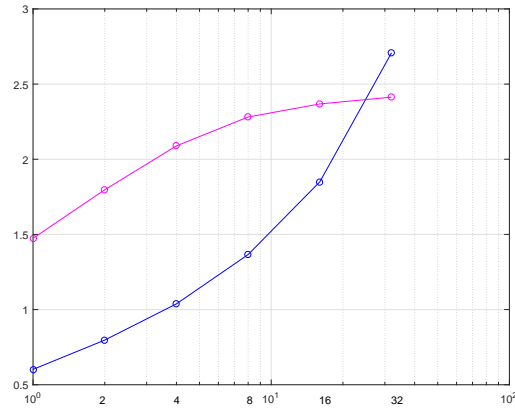




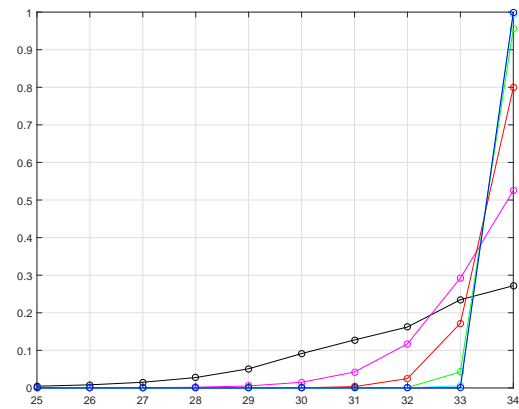
(a)



(b)



(c)



(d)

(a) free motion ( $w \equiv 0$ ) in  $(r, v)$ -plane; (b) Bayesian and worst-case risks of recovering  $w_t$  vs.  $t = 1, 2, \dots, 32$ ; (c) Bayesian and worst-case risks of recovering  $w^k := [w_{T-k+1}; w_{T-k+2}; \dots; w_T]$  vs.  $k$ ; (d) 10 largest eigenvalues  $\lambda_i(S_k)$  of  $S_k$  ( $k = 32, 16, 8, 4, 2$  and  $1$ ).

## Extension II: linear estimation over spectratopes

We say that a set  $\mathcal{X} \subset \mathbb{R}^n$  is a *basic spectratope*, if it can be represented in the form

$$\mathcal{X} = \{x \in \mathbb{R}^n : \exists t \in \mathcal{T} : R_k^2[x] \preceq t_k I_{d_k}, 1 \leq k \leq K\}$$

where

[S<sub>1</sub>]  $R_k[x] = \sum_{i=1}^n x_i R^{ki}$  are symmetric  $d_k \times d_k$  matrices linearly depending on  $x \in \mathbb{R}^n$  (i.e., “matrix coefficients”  $R^{ki}$  belong to  $\mathbf{S}^n$ )

[S<sub>2</sub>]  $\mathcal{T} \in \mathbb{R}_+^K$  is a convex compact subset of  $\mathbb{R}_+^K$  which contains a positive vector and is monotone:

$$0 \leq t' \leq t \in \mathcal{T} \Rightarrow t' \in \mathcal{T}.$$

[S<sub>3</sub>] Whenever  $x \neq 0$ , it holds  $R_k[x] \neq 0$  for at least one  $k \leq K$ .

A *spectratope* is a linear image  $\mathcal{Y} = PX$  of a basic spectratope.

We refer to  $d = \sum_k d_k$  as *size of the spectratope*  $\mathcal{Y}$ .

## Examples

[A.] Any ellitope is a spectratope.

[B.] Let  $L$  be a positive definite  $d \times d$  matrix. Then the “*matrix box*”

$$\mathcal{X} = \{X \in \mathbf{S}^d : -L \preceq X \preceq L\} = \{X \in \mathbf{S}^d : R^2[X] := [L^{-1/2}XL^{-1/2}]^2 \preceq I_d\}$$

is a basic spectratope. As a result, a *bounded* set  $\mathcal{X} \subset \mathbb{R}^\nu$  given by a system of “two-sided” LMI’s, specifically,

$$\mathcal{X} = \{x \in \mathbb{R}^\nu : \exists t \in \mathcal{T} : -t_k L_k \preceq S_k[x] \preceq t_l L_k, 1 \leq k \leq K\}$$

where  $S_k[x]$  are symmetric  $d_k \times d_k$  matrices linearly depending on  $x$ ,  $L_k \succ 0$  and  $\mathcal{T}$  satisfies  $S_2$ , is a basic spectratope:

$$\mathcal{X} = \{x \in \mathbb{R}^\nu : \exists t \in \mathcal{T} : R_k^2[x] \preceq t_k I_{d_k}, k \leq K\} \quad [R_k[x] = L_k^{-1/2} S_k[x] L_k^{-1/2}]$$

Same as ellitopes, *spectratopes admit fully algorithmic calculus.*

## Building linear estimate

**Proposition.** Consider convex optimization problem

$$\text{Opt} = \min_{H, \Lambda, \tau} \left\{ \begin{array}{l} \tau : (B - H^T A)^T (B - H^T A) \preceq \sum_k \mathcal{R}_k^*(\Lambda_k) + \tau S \\ \sigma^2 \text{Tr}(H^T H) + \phi_\tau(\lambda[\Lambda]) \leq \tau \end{array} \right\} \quad (*)$$

where  $\mathcal{R}_k^*(\Lambda) : \mathbf{S}^{d_k} \rightarrow \mathbf{S}^n$  is the conjugate linear mapping,

$$[\mathcal{R}_k^*(\Lambda)]_{ij} = \frac{1}{2} \text{Tr} (\Lambda [R^{ki} R^{kj} + R^{kj} R^{ki}]), \quad 1 \leq i, j \leq n,$$

and for  $\Lambda = \{\Lambda_k \in \mathbf{S}^{d_k}\}_{k \leq K}$ ,  $\lambda[\Lambda] = [\text{Tr}[\Lambda_1]; \dots; \text{Tr}[\Lambda_K]]$ .

Problem (\*) is solvable, and its feasible solution  $(H, \lambda, \tau)$  induces a linear estimate

$\hat{x}_H = H^T \omega$  of  $Bx$ ,  $x \in \mathcal{X}$ , via observation

$$\omega = Ax + \sigma \xi, \quad \xi \sim \mathcal{N}(0, I)$$

with  $S$ -risk not exceeding  $\sqrt{\tau}$ .

**Proposition.** Let  $\mathcal{X}$  be a spectratope, and let

$$\mathcal{Q} = \{Q \in \mathbf{S}_+^n : \exists t \in \mathcal{T} : \mathcal{R}_k[Q] \preceq t_k I_{d_k}, k \leq K\}, \quad \mathcal{Q}_\rho = \rho \mathcal{Q}, \rho > 0.$$

The set  $\mathcal{Q}$  is a nonempty convex compact set containing a neighbourhood of the origin, so that the quantity

$$M_* = \sqrt{\max_{Q \in \mathcal{Q}} \text{Tr}(BQB^T)},$$

is well defined and positive.

The efficiently computable linear estimate  $\hat{x}_{H_*}(\omega) = H_*^T \omega$  yielded by the optimal solution of (\*) is nearly optimal in terms of  $S$ -risk:

$$\text{RiskS}[\hat{x}_{H_*} | \mathcal{X}] \leq 2 \sqrt{2 \ln \left( \frac{8DM_*^2}{\text{RiskS}_{\text{Opt}}^2[\mathcal{X}]} \right)} \text{RiskS}_{\text{Opt}}[\mathcal{X}],$$

where

$$\text{RiskS}_{\text{Opt}}[\mathcal{X}] = \inf_{\hat{x}(\cdot)} \text{RiskS}[\hat{x} | \mathcal{X}]$$

is the minimax  $S$ -risk associated with  $\mathcal{X}$ , and  $d = \sum_k d_k$ .